Adjusted Predictions & Marginal Effects for Multiple Outcome Models & Commands (including ologit, mlogit, oglm, & gologit2)

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As was the case with logit models, the parameters for an ordered logit model and other multiple outcome models can be hard to interpret. Adjusted predictions and marginal effects can again make results more understandable. Stata 14 made the margins command much easier to use after multiple outcome commands like ologit, oprobit, mlogit, oglm and gologit2. While the examples here use ologit, the same procedures can be used with other commands. If you are condemned to using Stata 12 or 13, the appendix describes the much more tedious process that is required. However, regardless of which version of Stata you are using, you may prefer to use spost13 commands (findit spost13_ado) like mtable as they can produce easier to read output.

Consider the following. We will once again use the Nhanes2f data, but this time the dependent variable will be self-reported health, whose values range between 1 (poor) and 5 (excellent):

. webus . keep (2 obse . labe . labe . labe . labe . labe	if !mis ervation 1 define 1 define 1 values 1 values health	es2f, cl ssing(di ns delet e black	abetes (ed) 0 "nor 0 "ma black femal		"black"					
								Cum		
	1 poor 2 fair 3 aver 4 good 5 exce	2 2 2 age 3 2		729 1670 2938 2591	7.05 16.16 28.43 25.07 23.29		7.05 16.16 28.43 25.07 23.29	7.0 23.2 51.6 76.7 100.0)5 21 54 71	
Ordered	d logist	th i.fem tic regr d = -149	ressior		age, nol	og	LR ch Prob	i2(3)	=	10335 1682.10 0.0000 0.0534
ł	nealth	с С	Coef.	Std. Er	r.	z P	> z	[95% C	Conf.	Interval]
	female emale		70992	.035573	2 -3 .	29 0	.001	18682	215	0473769
		884		.058310 .001090						7702227 0389295
	/cut2 /cut3	-3.42	28162 4318	.074328 .064886 .058663 .056122	9 4			-3.5553 -2.1192	38 96	-4.765179 -3.300986 -1.88934 6412621

The results tell us that, on an all other things being equal basis, females, blacks, and older people tend to have lower levels of self-reported health. However, other than sign and significance, it is difficult to get a tangible feel for how large and important these differences are.

Using the Margins and Spost13 Commands. Now let's see what happens when we use the margins command to get the AAPs (Average Adjusted Predictions) for White people and Black people:

You get the AAPs by race for each category of the ordinal dependent variable. Looking at the results, you can see that, on an all other things equal basis, Black individuals are more than twice as likely as White individuals to say they are in poor health (13.7% as opposed to 6.4%), about as likely to report they are in average health (30.5% to 28.5%) and about half as likely to say they are in excellent health (12.3% versus 24.3%).

Long & Freese's mtable command (findit spost13_ado) produces output that is much easier to read and that does not require Stata 14:

You could simplify this a bit and say things like, according to the model, on an all other things being equal basis, almost 50% of White people say their health is good or excellent, compared to less than 31% of Black people. Or, if you prefer, you can say that more than 38% of Black individuals say their health is fair or poor, compared to less than 22% of White people. (Note that the AAPs for each group sum to 1, i.e. everybody has to fall into one of the five health categories.)

Given that there are a lot of numbers here, it might be easier to look at the AMEs:

Note: dy/dx for factor levels is the discrete change from the base level.

. mtable, dydx(black)

Expression: Marginal effect of Pr(health), predict(outcome())

poor	fair	average	good	excellent
0.074	0.093	0.021	-0.068	-0.120

Consistent with the earlier results, the marginal effects show you that, on average, Black individuals are 7.4 percentage points more likely than White people to say their health is poor, and about 12 percentage points less likely to say their health is excellent. Personally, I find numbers like this much more tangible and meaningful than the raw coefficients.

Here is an example of how you can use Long & Freese's mtable command to compute adjusted predictions for prototypical cases. The first command does everything with one command but the subsequent commands produce an easier to read table.

```
. * mtable
```

. mtable, at (black = (0 1) age = 20) at (black = (0 1) age = 47) at (black = (0 1) age = 74) dec(4)

Expression: Pr(health), predict(outcome())

d	lack age	e poor	fair	average	good	excellent
1 2 3 4 5 6	0 20 1 20 0 4 ² 1 4 ² 0 7 ⁴ 1 7 ⁴	0.0414 70.0513 70.1157 40.1407	0.0553 0.1184 0.1410 0.2498 0.2781 0.3517	0.1731 0.2813 0.3046 0.3395 0.3305 0.2430	0.2869 0.2930 0.2786 0.1882 0.1634 0.0834	0.4672 0.2659 0.2245 0.1068 0.0872 0.0380
 quietly mtable, quietly mtable, quietly mtable, quietly mtable, quietly mtable, mtable, at (black) 	at (black =) at (black =) at (black =) at (black =)	L age = 20)) age = 47) L age = 47)) age = 74)	rown (20 ye rown (47 ye rown (47 ye rown (74 ye	ar old blac ar old white ar old blac ar old white	k) dec(4) e) dec(4) k) dec(4) e) dec(4)) below) below) below

Expression: Pr(health), predict(outcome())

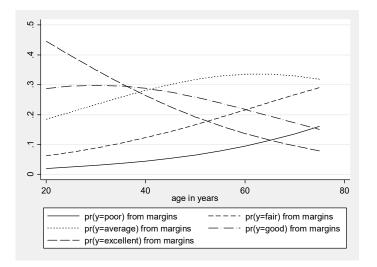
	poor	fair	average	good	excellent
	1		0 1701		0 4670
20 year old white	0.0175	0.0553	0.1731	0.2869	0.4672
20 year old black	0.0414	0.1184	0.2813	0.2930	0.2659
47 year old white	0.0513	0.1410	0.3046	0.2786	0.2245
47 year old black	0.1157	0.2498	0.3395	0.1882	0.1068
74 year old white	0.1407	0.2781	0.3305	0.1634	0.0872
74 year old black	0.2839	0.3517	0.2430	0.0834	0.0380

Specified values of covariates

	Ι	black	age
	+-		
Set 1	Ι	0	20
Set 2	Ι	1	20
Set 3	Ι	0	47
Set 4		1	47
Set 5		0	74
Current	Ι	1	74

Producing readable graphs can be a little tricky since so many different lines can be produced at the same time because of the multiple outcomes for the dependent variable. The spost13 mgen command might be useful for this purpose. I'll just show the commands and then the final graph.

```
* Graphics using mgen
* mgen for all groups pooled together
mgen, at(age = (20(5)75)) stub(all)
list allpr1 allpr2 allpr3 allpr4 allpr5 allage in 1/15
line allpr1 allpr2 allpr3 allpr4 allpr5 allage, scheme(sj) name(pooled)
```



The line that starts at the top left shows that the probability of being in excellent health is around 45% for twenty year olds and less than 10% by age 70. Conversely, the line that starts at the bottom left shows that hardly anyone is in poor health at age 20 but by age 70 close to 20% are.

If you are doing multiple groups (e.g. Black individuals and White individuals) you may want to graph only some of the outcomes. In the following graph we include only the predicted probabilities of poor health and excellent health, separately for Black people and White people.

```
* mgen for groups
drop allpr1 - allCpr5
mgen, at(age = (20(5)75) black = 0) stub(wh) predn(whpr)
mgen, at(age = (20(5)75) black = 1) stub(bl) predn(blpr)
line whwhpr1 blblpr1 whwhpr5 blblpr5 whaqe, scheme(sj) name(byrace)
  ß
  4
  arepsilon
  2
  0
     20
                      40
                                        60
                                                         80
                            age in vears
            whpr(y=poor) from margins
                                       blpr(y=poor) from margins
           whpr(y=excellent) from margins
                                       blpr(y=excellent) from margins
```

If you look at the two lines that start at the bottom left, you see that the likelihood of being in poor health increases with age and that at every age Black individuals are more likely to be in poor health than are White people (and the gap increases with age). Conversely the top two lines

show you that the probability of being in excellent health decreases with age and that at every age Black people are less likely to have excellent health than are White people.

You may also want to look at the user-written combomarginsplot, available from SSC. It shows how you can combine categories (e.g. poor and fair, good and excellent) when doing a graph.

Finally the spost13 mchange command can be good for producing lots of potentially useful statistics. Read the help for the command if you aren't already familiar with it.

```
. * mchange
```

```
. mchange black female age, stats(change start end) dec(5) delta(10)
```

ologit: Changes in Pr(y) | Number of obs = 10335

Expression: Pr(health), predict(outcome())

	poor	fair	average	good	excellent
black					
black vs nonBlack	0.07385	0.09265	0.02085	-0.06771	-0.11964
From	0.06358	0.15488	0.28464	0.25405	0.24284
To	0.13743	0.24753	0.30549	0.18634	0.12321
female					
female vs male	0.00743	0.01179	0.00652	-0.00681	-0.01893
From	0.06707	0.15829	0.28411	0.25087	0.23966
To	0.07450	0.17008	0.29062	0.24406	0.22074
age					
+1	0.00266	0.00415	0.00220	-0.00244	-0.00657
From	0.07100	0.16450	0.28747	0.24726	0.22977
To	0.07365	0.16865	0.28967	0.24482	0.22321
+delta	0.03077	0.04251	0.01537	-0.02859	-0.06006
From	0.07100	0.16450	0.28747	0.24726	0.22977
To	0.10177	0.20701	0.30284	0.21867	0.16972
Marginal	0.00261	0.00413	0.00227	-0.00239	-0.00663
From	• Z	• Z	• Z	• Z	• Z
To	• Z	• Z	• Z	• Z	• Z

Average predictions

	poor	fair	average	good	excellent
Pr(y base)	0.07100	0.16450	0.28747	0.24726	0.22977

Appendix: Adjusted Predictions & Marginal Effects for Multiple Outcome Models & Commands (Stata 13 and Earlier)

NOTE: Stata 14 made it much easier to estimate adjusted predictions and marginal effects for multiple outcome commands like mlogit and ologit and oglm and gologit2. Therefore use Stata 14 or higher, or the spost13 commands, if at all possible. However if you are condemned to using Stata 12 or 13 and do not want to use spost13, this appendix shows the (much more tedious) commands that are needed. Otherwise you can skip this appendix. ologit is used in these examples but similar and probably identical procedures can be used with other multiple outcome commands.

As was the case with logit models, the parameters for an ordered logit model can be hard to interpret. Adjusted predictions and marginal effects can again make results more understandable. The margins command, however, is harder to use, because we need to issue a separate command for *each* of the outcomes of the ordinal variable. Consider the following. We will once again use the Nhanes2f data, but this time the dependent variable will be self-reported health, whose values range between 1 (poor) and 5 (excellent):

. webus . keep (2 obse . label . label . label . label . fre h	<pre>version 13.1 webuse nhanes2f, clear keep if !missing(diabetes, black, female, age) (2 observations deleted) label define black 0 "nonBlack" 1 "black" label define female 0 "male" 1 "female" label values black black label values female female fre health health 1=poor,, 5=excellent</pre>									
					 Valid					
Valid	1 poor 2 fair 3 aver 4 good 5 exce		729 1670 2938 2591 2407	7.05 16.16 28.43 25.07 23.29	28.43 25.07 23.29	7.05 23.21 51.64 76.71 100.00				
Ordered	d logist	<pre>ch i.female i cic regression d = -14923.34</pre>	n		Numbe LR ch Prob	r of obs = i2(3) = > chi2 = o R2 =	0.0000			
	nealth				₽> z	[95% Conf.	Interval]			
	female emale		.0355732	-3.2	9 0.001	1868215	0473769			
k						998796 043205				
	/cut2	-4.910859 -3.428162 -2.004318 7512595	.0648869			-5.056539 -3.555338 -2.119296 8612569	-3.300986			

The results tell us that, on an all other things being equal basis, females, Black people, and older people tend to have lower levels of self-reported health. However, other than sign and significance, it is difficult to get a tangible feel for how large and important these differences are.

Using the Margins Command. Now let's see what happens when we use the margins command to get the AAPs (Average Adjusted Predictions) for White individuals and Black individuals:

. margins blac	ck					
Predictive man Model VCE	Number	c of obs =	10335			
<pre>Expression : Pr(health==1), predict()</pre>						
		Std. Err.	Z	₽> z	[95% Conf.	Interval]
black nonBlack black	.0635826 .1374343	.0023288 .0070777	27.30 19.42	0.000 0.000	.0590182 .1235622	.068147 .1513064

We see that Black people score more than 7 percentage points higher than do White people – which may seem odd given that the coefficient for black is negative. However, you have to realize that this is ONLY the predicted probability for health = 1, i.e. the probability that somebody will say their health is poor. So, it is not surprising that blacks are more than twice as likely to report they are in poor health as whites are. This is only one of the five possible outcomes though; if we want to get all of them we need to run five separate margins commands.

. margins black, predict(outcome(#2)) Number of obs = 10335 Predictive margins Model VCE : OIM Expression : Pr(health==2), predict(outcome(#2)) _____ Delta-method Margin Std. Err. z P>|z| [95% Conf. Interval] ____+ _____ black | nonBlack | **.1548809** .0034868 44.42 0.000 .1480469 .161715 black | **.2475271** .0072026 34.37 0.000 .2334102 .261644 . margins black, predict(outcome(#3)) Predictive margins Number of obs = 10335 Model VCE : OIM Expression : Pr(health==3), predict(outcome(#3)) _____ Delta-method z P>|z| [95% Conf. Interval] Margin Std. Err. -----black | nonBlack | .2846398 .0043742 65.07 0.000 .2760666 .293213 black | .3054912 .0048495 62.99 0.000 .2959863 .3149961 _____ . margins black, predict(outcome(#4)) Number of obs = 10335 Predictive margins Model VCE : OIM Expression : Pr(health==4), predict(outcome(#4)) _____ 1 Delta-method i | z P>|z| [95% Conf. Interval] Margin Std. Err. black | .1750961 .2624894 nonBlack.2540548.004303559.030.000.2456201black.1863406.005737132.480.000.1750961 _____ . margins black, predict(outcome(#5)) Number of obs = 10335 Predictive margins Model VCE : OIM Expression : Pr(health==5), predict(outcome(#5)) _____ Delta-method Margin Std. Err. z P>|z| [95% Conf. Interval] 1 black | nonBlack | **.2428419** .0041509 58.50 0.000 .2347062 .2509776 black | **.1232068** .0060262 20.45 0.000 .1113957 .135018

Looking at the results, you can see that, on an all other things equal basis, Black individuals are more than twice as likely as White individuals to say they are in poor health (13.7% as opposed to 6.4%), about as likely to report they are in average health (30.5% to 28.5%) and about half as likely to say they are in excellent health (12.3% versus 24.3%). You could simplify this a bit and

say things like, according to the model, on an all other things being equal basis, almost 50% of whites say their health is good or excellent, compared to less than 31% of blacks. Or, if you prefer, you can say that more than 38% of Black individuals say their health is fair or poor, compared to less than 22% of non-black people. (Note that the AAPs for each group sum to 1, i.e. everybody has to fall into one of the five health categories.)

You also get the marginal effects one outcome at a time. For variety, rather than use the dydx option, I will use the r. operator, which gives you some additional useful output. For convenience, I will just show the two most extreme outcomes, poor health (outcome 1) and excellent health (outcome 5).

. * AMEs using margins . margins r.black, predict(outcome(#1))									
Contrasts of predictive margins Model VCE : OIM									
<pre>Expression : Pr(health==1), predict(outcome(#1))</pre>									
	df	chi2	P>chi2						
black	1 13	81.83	0.0000						
	Contrast	Std.	ethod Err. 	[95% Conf.	Interval]				
black (black vs nonBlack)		.0738517 .006432			.0864582				
. margins r.black, pre	dict(outcom	ue (#5))							
Contrasts of predictiv Model VCE : OIM	ve margins								
Expression : Pr(heal	Lth==5), pre	dict(out	.come(#5))					
	df	chi2	P>chi2						
	1 34	1.63	0.0000						
	Delta-method Contrast Std. Err.								
black (black)									

Consistent with the earlier results, the marginal effects show you that, on average, Black individuals are 7.4 percentage points more like than White individuals to say their health is poor, and about 12 percentage points less likely to say their health is very good.

Personally, I find numbers like this much more tangible and meaningful than the raw coefficients; but the process of generating them is tedious and prone to error. It is probably even worse when you want to make statements about the effects of continuous variables. Fortunately, the m* commands in Spost13 and other user-written routines can make these tasks easier.