ORIGINAL PAPER

Combining physical-based models and satellite images

3 for the spatio-temporal assessment of soil infiltration capacity

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Abstract The performance of managed artificial recharge (MAR) facilities by means of surface ponds (SP) is controlled by the temporal evolution of the global infiltration capacity I_c of topsoils. Cost-effective maintenance operations that aim to maintain controlled infiltration values during the activity of the SP require the full knowledge of the spatio-temporal variability of I_c . This task is deemed uncertain. The natural reduction in time of I_c depends on complex physical, biological and chemical reactions that clog the soil pores and has been observed to decay exponentially to an asymptotic non-zero value. Moreover, the relative influence of single clogging processes depend on some initial parameters of the soil, such as the initial infiltration capacity $(I_{c,0})$. This property is also uncertain, as aquifers are typically heterogeneous and scarcely characterized in practical situations. We suggest a method to obtain maps of I_c using a geostatistical approach, which is suitable to be extended to engineering risk assessment concerning management of SP facilities. We propose to combine geostatistical inference and a temporally-lumped

topsoils of a SP, using field campaigns of local and large scale tests and additionally by means of satellite images as secondary information. We then postulate a power-law relationship between the parameter of the exponential law, λ , and $I_{c,0}$. It is found that calibrating the two parameters of the power law model it is possible to fit the temporal evolution of total infiltration rate at the pond scale in a MAR test facility. The results can be used to design appropriate measures to selectively limit clogging during operation, extending the life of the infiltration pond.

Keywords Managed artificial recharge · Surface ponds · Infiltration capacity · Clogging · Collocated cokriging · Satellite images

1 Introduction

Managed artificial recharge (MAR) practices include several methods that aim to recover and enhance groundwater quality and productivity of depleted aquifers (Bouwer 2002). Among them, the use of excavated artificial ponds, or surface ponds (SP) are widely common (e.g. Management of aquifer and subsurface (2003), Scanlon et al. (2006)).

In an artificial SP facility, recharge is induced by flooding the excavation with water coming from any available source (e.g. reclaimed water, stormwater, river water), which percolates towards the subsurface by infiltration. In a properly designed facility, evaporation and other losses are typically negligible (Bouwer 1999) compared with the infiltration rates. The maximum rate at which water can infiltrate in the subsurface is known as the infiltration capacity (I_c). It regulates both the total amount of infiltration towards the aquifer and some characteristic

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times (such as the water residence time within an SP) which are needed to make appropriate management of the facilities (Bouwer 2002; Pedretti et al. under review; Perez-Paricio and Carrera 1999).

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Prediction of I_c can be done at two characteristic spatial scales. At the pond scale, from a purely managerial perspective, the spatial average infiltration in the SP, $\bar{I}_c(t)$, provides quantitative information about the total amount of actual infiltration in an integrated way (Barahona-Palomo et al. 2010). Monitoring the SP facility during operation allows the evaluation of the integrated infiltration capacity with time. However, this does not provide any information about the spatio-temporal distribution of the local scale infiltration capacity $(I_c(\mathbf{x},t))$, which is crucial for the optimal operation, management, and maintenance (O& M) of the artificial SP facility (Pedretti et al. under review).

Assessing the full spatio-temporal distribution of $I_c(\mathbf{x}, t)$ during operation is not realistically feasible. During recharge processes, I_c reduces in time because of a variety of processes that tend to modify the pore size distribution (PSD) of the ground, namely reducing the porosity of the top soil (e.g. Bouwer 2002). The topsoil hydraulic conductivity $K_s(\mathbf{x},t)$ depends on the PSD, which in turn controls $I_c(\mathbf{x},t)$. To complicate matters even further, natural "pore clogging" also depends on some initial properties of the soil (e.g. Guin 1972). In typical situations, the impossibility of complete characterization of geological heterogeneities make these variables random parameters for modelling purposes (Tartakovsky and Winter 2008). In this sense, a probabilistic approach is advisable (Krzysztofowicz 2001), since it could then be integrated into a larger scale probabilistic risk assessment studies (Tartakovsky 2007; Bolster et al. 2009; de Barros and Rubin 2008; Masetti et al. 2009).

In this work we explicitly exclude accidental extreme events (with very low frequency/probability) that suddenly reduce global infiltration (e.g. Bedford and Cook 2001, pp. 103-104). Natural reactions, on the other hand, are more frequent and are related to a variety of physical, biological and chemical mechanisms (e.g. Baveye et al. 1998; Vandevivere et al. 1995; Greskowiak et al. 2005) that together reduce I_c during artificial recharge. As a consequence, the rate of natural clogging mechanisms at small scales is highly uncertain. Several approaches have been suggested in the past to assess the effect of independent individual mechanisms. For instance, the soil can be conceptualized as a natural filter in which suspended particles circulate in the soil while percolating; they are attracted and collapse into the pore spaces. At some macroscopic scale, this effect is called "physical clogging" and has been modeled using "depth filter" theories (see Zamani and Maini 2009; Pedretti et al. under review; Guin 1972 for details). A significant point is that the capacity of the soil to retain the particles (i.e. the potential of the soil to be clogged) depends on whether surface forces or volumetric forces dominate. Bioclogging can also be modeled using a macroscopic approach considering the expansion of microbial colonies to create a biomass that get trapped because of similar forces that affect non-organic particles (Baveye et al. 1998; Vandevivere et al. 1995; Clement et al. 1996).

The big question, however, remains how to evaluate the effect on pore clogging that encompasses the mutual dependence of all clogging mechanisms. Although several published field and laboratory experiments exist that aim to understand and quantify their mutual interaction depending on soil type (Perez-Paricio and Carrera 1999; Kim et al. 2010; Hoffmann and Gunkel 2010), predicting the clogging rates via cumulation of individual mechanisms remains, in the practice, challenging. The reasons are that : (i) reactions leading to a (macroscale) clogging effect take place at microscales (e.g. Baveye et al. 1998) that cannot be easily measured, or that cannot be easily upscaled to the field pond scale and (ii) the rate at which clogging mechanisms jointly develop depends on a variety of site-specific conditions and factors, such as textural heterogeneities of the soil (e.g. pore or grain size distributions—PSD/GSD) (Zamani and Maini 2009; Tien and Payatakes 1979; Guin 1972), soil density heterogeneities (Clement et al. 1996), chemical heterogeneity (Greskowiak et al. 2005), and fluctuations of the water temperature (Civan 2007, 2010).

The quality of most of the information depends on the quality of geological characterizations, which is almost always poor in practice since direct investigations are costly, affected by serious errors and sometimes simply not feasible. Some models have been proposed in the past to correlate initial hydraulic conductivity profiles with the intensity of clogging. Unfortunately, most of them are based on assumptions of homogeneities that over- or underestimate the rate of clogging. For instance, assuming the soil pore sizes or the grain sizes to be uniformly (e.g. Kozeny 1927; Carman 1938; Hazen 1882) and non-uniformly (e.g. Guin 1972) distributed is essential to correctly estimate the development of clogging. Indeed, the use of uniform formulations implicitly include average macrocharacteristic features of PSD and GSD, such as the total soil porosity ϕ or some characteristic grain size d_{ϱ} , from which the seepage velocity is calculated. However, bioclogging does not grow uniformly but rather starts by developing local microcolonies within smaller pores, and grows to occupy the larger pores. On the other hand, physical mechanisms act differently. According to the filter theory (see Zamani and Maini 2009; Pedretti et al. under review; Guin 1972 for details) the clogging rates are inversely proportional with d_g but this relationship depends strongly on other factors such as the suspended solid size (d_s) and $I_c(t=0)$. Guin (1972) adopted a macroscopic

Kozeny-like model where the porous medium is decomposed into a bundle of channels with different hydraulic radii R (ratio between porosity and the specific surface area of the soil, S_a). In the case of rapid particle deposition, the clogging rate of individual pores is proportional to the square of its specific surface area $\left(-\frac{\mathrm{d}S_a}{\mathrm{d}t} \propto S_a^2\right)$. Thus, this implies that the area of larger pores (i.e. larger initial infiltration capacities) is decreased preferentially. In the case of slow particle deposition, $-\frac{\mathrm{d}S_a}{\mathrm{d}t} \propto S_a^{1/2}$. Experiments on bioclogging (Cunningham et al. 1991; Vandevivere and Baveye 1992) showed that the relative change of permeability (K_s/K_{s0}) depends, among other parameters, on the GSD.

It has been observed that all clogging mechanisms do not go on indefinitely, but rather end up providing an asymptotic infiltration value, which varies in space (e.g. Baveye et al. 1998). The reasons that lead to an asymptotic clogging effect are controversial and depend on local conditions. For instance, bioclogging mechanisms roughly follow a Monod behavior (Monod et al. 1965; Matsumoto 1979; Baveye et al. 1998), which reproduce a microbial growth with a maximum asymptotic value of development. Physical clogging, on the other hand, varies in time since the governing forces change from volumetric to surface and vice-versa depending on flow velocity, available reactive surface areas, etc. (see Zamani and Maini 2009 for a complete description of whole physical processes). While physical clogging could theoretically result in zero infiltration, bioclogging may help preventing additional physical clogging, so that the combination of both processes may still present a finite non-zero asymptotic value.

While assessment of initial and final permeabilities can certainly be done with the use of empirical or theoretical formulations, the question still remains about how to evaluate the actual evolution of infiltration with time. This is required for a proper risk assessment (Pedretti et al. under review). Lumped solutions such as exponential decay formulations are often adopted at the field scale (Iwasaki 1937; Perez-Paricio and Carrera 1999; Kim et al. 2010; Hoffmann and Gunkel 2010). The use of such models is appealing since only a few parameters are required for estimation purposes, but so far there is no widely accepted work on the relationship between the upscaled and the local infiltration models. While lumped (upscaled) models are routinely used as a way to assess when maintenance operations should take place at the full pond (usually drying the pond and cleaning), we contend that a detailed clogging model could be used to derive an ad-hoc remediation operation that targets only specific portions of the pond (similar to the concept of precise agriculture).

A major limitation of mapping infiltration variations locally in space and time is that primary information is

limited and plagued by errors. A viable option to enhance the mapping of characteristic soil properties is to combine direct (primary) with secondary data (related to the primary ones). An example is to combine measurements from surface infiltrometers with satellite-images-based secondary information (Pedretti et al. 2010). Normally, secondary information is suitable for spatial assessment as it typically contains a denser dataset (e.g. Gooverts 1997), but the robustness of the method relies in the strength of the correlation existing between the two types of information.

In this paper, we develop a geostatistical approach to map soil clogging parameters integrating a limited dataset of direct information with secondary information based on satellite images. A physically-based model is used to evaluate the spatial-temporal decreasing of the local infiltration capacities throughout the SP.

The paper is structured as follows. In Sect. 2 we propose a local infiltration model with an exponential decay to an asymptotic value. Initial, asymptotic infiltration, and decay coefficients are considered to be correlated. In Sect. 3 we further develop the model by applying it directly to an artificial infiltration pond where local infiltration data is available at a few points and can be further obtained from a full mapping of a secondary variable. An external validation of the model from the evolution of the spatially averaged infiltration with time is also included. This is followed by the main conclusions of the work.

2 Methodology

2.1 Problem statement

Let us consider a given artificial SP recharge site in which direct measurements of the local infiltration rate, $I_c(\mathbf{x},t)$, are known at sparse locations and at a few discrete times, i.e., $I_c(\mathbf{x}_i,t_j)\{i=1,\ldots,m\},\{j=1,\ldots,n\}$. This is often the case in most practical applications as the operation of an artificial recharge pond is typically done under flooding conditions and thereby any exhaustive characterization of the infiltration capacity is too costly and time-consuming. The support scale of the measurement is local and given, for example, by the size of an infiltrometer test.

We assume that an extensively sampled secondary variable is available at some support scale. An example would be data coming from the colour intensity of a satellite image, which can provide valuable information related to the soil hydraulic properties such as the moisture content of the soil, the vegetative canopy density on the ground (Chica-Olmo and Abarca-Hernandez 2000; Granger 2000; Milewskia et al. 2009) or the infiltration capacity



at a specific time (Pedretti et al. 2010). Evidently, a limitation of satellite images is that image-based methods are restricted to non-flooded times, with at most two measuring times, i.e., before and after flooding.

Under these conditions, we attempt to map the temporal evolution of the infiltration capacity at the site so that better management of the artificial recharge pond can be undertaken.

2.2 Clogging model

We assume that the temporal reduction of the infiltration capacity at every point **x** in the domain due to clogging processes follows a decaying exponential law. The model assumes that clogging takes place on the order of tens of days. Temperature can therefore be neglected as it fluctuates on two temporal scales, seasonally and daily, that do not affect clogging occurrence. Seasonality occurs at temporal scales much longer than clogging time, and night-day fluctuations are too short to affect clogging development permanently. The clogging model can be formulated in different ways. One possible formulation is

$$I_c(\mathbf{x}, t) = RI_c(\mathbf{x}, t = 0) \exp(-\lambda(\mathbf{x}, t)t)$$

= $I_{c,0}(\mathbf{x}) \exp(-\lambda(\mathbf{x}, t)t),$ (1)

where $I_c(\mathbf{x}, t = 0)$ is the infiltration capacity at some initial operation stage, $\lambda(x, t)$ is the lumped clogging coefficient and R is a generic, instantaneous reduction function of infiltration caused by gas production and other mechanisms (Olsthoorn 1982; Bouwer 2002), which in general is close to 1. We can remove the impact of R by using $I_{c,0}(\mathbf{x}) = R$ $I_c(\mathbf{x}, t = 0)$ as the initial infiltration value. It is worth noting that in (1), λ changes over space and time, tending to zero as time increases. Consequently, $I_c(\mathbf{x}, t)$ values tend to an (spatially dependent) asymptotic value, $I_{c,f}$

Therefore, an alternative model to (1) that implicitly includes this asymptotic value $I_{c,f}$ can be formulated as

$$I_c(\mathbf{x},t) = I_{c,f}(\mathbf{x}) + (I_{c,0}(\mathbf{x}) - I_{c,f}(\mathbf{x})) \exp(-\lambda(\mathbf{x})t), \qquad (2)$$

where $\lambda(\mathbf{x})$ is now constant in time. This clogging model constitutes the basis of our approach to map the temporal evolution of the infiltration capacity. The fundamental advantage of this model is that parameterizes the temporal evolution of the infiltration capacity I_c by means of three constant-in-time variables/parameters: $I_{c,f}(\mathbf{x})$, $I_{c,0}(\mathbf{x})$, and $\lambda(\mathbf{x})$, all three variable in space, and are therefore susceptible to simple geostatistical analysis.

Total infiltration capacity can be obtained by spatial averaging of either (1) or (2); i.e.

$$\overline{I}_c(t) = \frac{1}{V} \int_V I_c(\mathbf{x}, t) d\mathbf{x}.$$
 (3)

In most cases $\overline{I}_c(t \to \infty) \equiv \overline{I}_{cf}$ will be too small to be acceptable; that is, it will be below a pre-specified threshold value, I_t . Such small could be unacceptable in practical situations. Thus, the need for a model that can provide information about the spatial evolution of infiltration with time. It turns out that whenever $I_{c,f}(\mathbf{x})$ and $I_{c,0}(\mathbf{x})$ are fully known, the temporal evolution of the local infiltration depend exclusively on λ (\mathbf{x}). We postulate that this non-time dependent clogging factor is directly correlated with some initial property of the soil, so that λ (\mathbf{x}) = $f(I_{c,0}(\mathbf{x}))$. This is feasible since clogging develops at different rates according with the local distribution of PSD or GSD, both highly uncertain.

2.3 Mapping the soil infiltration capacity

If experimental measurements of I_c are limited to a few locations, secondary information can be incorporated to estimate the spatial distribution of the primary variable. For this purpose, several geostatistical techniques can be used, such as the collocated cokriging model. Yet, in this case, its direct application is cumbersome as one needs to estimate the evolution of the variogram matrix with time as clogging progresses. To overcome this problem, we propose a new approach.

We start by noting that Pedretti et al. (2010) observed that $I_{c,0} = f_1(P_v)$ and $I_{c,f} = f_2(P_v)$, where P_v are the color pixel values of a satellite image. By extension, and invoking a phenomenological approach, it is to be expected that there could exist a relationship between soil parameters and P_v , and thus also between λ and P_v .

This relationship should be constructed via theoretical or empirical methods. The theory suggests that some clogging mechanisms, such as physical clogging, can be modeled using a filter approach (e.g. Zamani and Maini 2009; Pedretti et al. under review); thus, the clogging rate should be negatively correlated with some grain size representative diameter, d_{g} . Other mechanisms however do not necessarily rely on this assumption: biological clogging can be modeled using Monod-based growing models (e.g. Clement et al. 1996), for which the rate of bioclogging is linearly proportional to the soil density ρ which is positively correlated with d_{ϱ} (Pedretti et al. under review). Therefore, in a real site, it is important to assess the relative importance of the two mechanisms, since this will control the relationship existing between λ and the soil properties. In the field though, the use of (2) gives lumped clogging factors, in which the single effects of each mechanism is somehow hidden. As such, any relationship existing between P_{ν} and experimentally-based λ does not give any indication of the relationship existing between P_{ν} and all the parameters characterizing the soil, including d_g , but

also density, fraction of organic matter, etc. Thus, it is not a priori clear whether λ and P_{ν} should be positively or negatively correlated. In the application example later we will use both possibilities.

Our approach starts from having a perfect knowledge of some secondary variable (obtained for example from satellite images), while there is little to no information on the three primary variables controlling infiltration: $I_{c,0}$, $I_{c,f}$, and λ . Thus, it is possible to use some geostatistical approach involving either cokriging (to get a smooth estimate in a mean sense) or cosimulation (to be included in a Monte Carlo approach). Since data from the primary variables are always expected to be less than exhaustive, we chose a collocated cokriging approach under the Markov model I. Such a model limits the secondary variable to the data available at the estimation location, and further estimates the cross-variograms by employing an underlying regression model. The advantage is that matrix instabilities caused by densely sampled secondary data (such as high-resolution satellite rasters of pixel values) are avoided and reduce the burden of modelling all variograms and cross-variograms in a cokriging system (Almeida and Journel 1996; Journel 1999). In this case, the cokriging model only requires the knowledge of the variogram of the primary variable, the correlation coefficient, and the variance of the secondary variable. Thus, the modelling effort is almost the same as for kriging one variable. For simplicity, the cokriging of each variable is conducted independently.

Unfortunately, in most practical situations, the variogram of these primary variables cannot be directly estimated as too little information is available at a given SP site. To circumvent this problem, we suggest to completely rely on the variogram of the secondary variable (satellite image) that is always well characterized. For each primary variable, $\{I_{c,0}, I_{c,f}, \lambda\}$, the auto-variogram can be estimated using regression models. Pedretti and others (Pedretti et al. 2010) found that a linear regression satisfactorily correlates the natural logarithm of infiltration capacities (both $Y_0 = \ln(I_{c,0})$) and $Y_f = \ln(I_{c,f})$) with the color pixel values of an image (P_v) . The model has the form

$$Y_0 = a_0(P_v) + b_0 + \varepsilon_{Y,0},\tag{4}$$

$$403 Y_f = a_f(P_v) + b_f + \varepsilon_{Yf}, (5)$$

where a and b are the regression coefficients, and ε_Y represents regression model errors. To complete the picture we must specify a model for λ . Parsimony leads us to

408 postulate a similar linear model in terms of $Y_{\lambda} = \ln \lambda$,

$$Y_{\lambda} = a_{\lambda}(P_{\nu}) + b_{\lambda} + \varepsilon_{Y,\lambda}. \tag{6}$$

From these simple models we can write the variogram functions as

$$\gamma_{Y_i}(\mathbf{h}) = a_i^2 \gamma_{P_v}(\mathbf{h}) + \gamma_{\varepsilon_{Y_i}}(\mathbf{h}), \quad i = 0, f, \lambda$$
 (7)

where \mathbf{h} is the lag distance between data values. We can further assume that the regression model errors are uncorrelated both with errors at different locations or with the secondary variable (pure nugget).

Once the spatial distributions of the three variables have been obtained, it is possible to use (2) to obtain the spatiotemporal distribution of the local infiltration rate, and consequently the temporally variable global infiltration. We now illustrate the method and highlight the limitations in a real site.

3 Application example

A pilot SP is located in Sant Vicenç dels Horts, close to the city of Barcelona (Spain). The geological background is a sequence of deltaic deposits of the river Llobregat. A high resolution image was captured in November 15, 2007 (Fig. 1), representing the ground conditions at the site before a flooding experiment was conducted. We will consider this as the initial state of our system, from which we wish to obtain a map of $I_{c,0}$ values. The resolution of the image is around 0.5 m² per pixel. The infiltration area is approximately 100 m x 46 m. The unconfined and highly transmissive aquifer below the artificial pond is between 12 and 15 m in thickness, and has suffered from overexploitation and pollution in recent years due to an increasing population and industrial density in the Barcelona area (Carrera et al. 2005; Custodio 2002).

3.1 Observations and modelling at the pond scale (global scale)

A flooding test was performed between March and June 2009 in order to test the performance of the site for MAR operations. The following variables were recorded on a continuous basis: headwater at the pond, h(t); discharge rate towards the aquifer, $Q_{IN}(t)$ (recorded from water mass balance within the pond, disregarding evaporation); infiltration area, A(t) (usually a direct function of h(t) to include the basin slopes); and distance from the surface to the water table, L(t). Infiltration at the full pond scale is then obtained as

$$\overline{I}_c(t) \approx \left(\frac{Q_{IN}(t)L(t)}{Ah(t)}\right)$$
 (8)

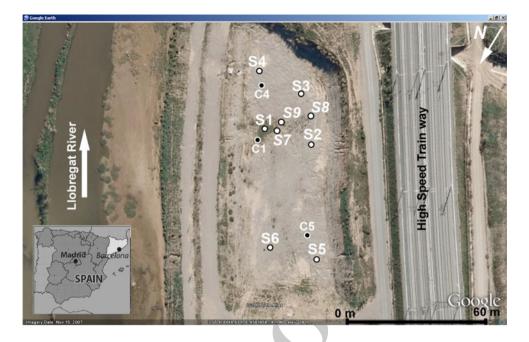
The recorded infiltration values can be matched by a simple exponential model

$$\overline{I}_c(t) = \overline{I}_{c,f} + (\overline{I}_{c,0} - \overline{I}_{c,f}) \exp(-\lambda^e t)$$
(9)

Figure 2 shows that a good match between data and the best-fit approximation of the large-scale infiltration model 456



Fig. 1 Satellite image of the Sant Vicenç dels Horts EAP, close to Barcelona (property of Google and ICC, 2007). S-tags refer to the locations of infiltrometer tests; C-tags are excavated pits to observe the local geological stratigraphy



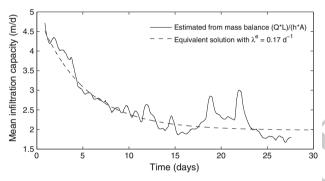


Fig. 2 Experimental observations (*solid line*) of $\overline{I_c}$ with time from the flooding test. The *dotted curve* is the best-fit approximation with $\overline{I}(t=\infty)=1.98$ m,/d and $\lambda^e=0.17$ day⁻¹

(9) can be obtained. λ^e is an equivalent clogging parameter for the entire pond. Note that the asymptotic value is used to evaluate the minimum infiltration value of the pond, which is reached when clogging no longer develops. The solid line represents observations using (8), while the dotted line is the best fit model solution using (9) with the following parameters: $\overline{I}_{c,0} = 5.35 \text{m/d}$, $\overline{I}_{c,f} = 1.98 \text{ m/d}$, and $\lambda^e = 0.17 \text{ day}^{-1}$.

Note that according to the calibration process using (9), we have a global model of the reduction of infiltration with time, that could now be used for management or risk evaluations. For example, it is found that reduction of infiltration to 50% of the initial value is reached after just roughly 7 days, but it will for example take 14 days to reduce to 40% of the original value. It is possible to obtain a compromise between allowing the system to work for larger periods without maintenance by just allowing less water to infiltrate.

In Fig. 2 the infiltration curve displays some daily oscillations. This is due to unaccounted causes, such as the impact of water temperature or atmospheric pressure with time. A detailed analysis of these causes could lead to a smoothing of the oscillations, leading to a better fit of the simple exponential model. However, this is not deemed essential here and the impact of these additional processes is not pursued in our work.

3.2 Observations at the local scale

Double-ring infiltrometers (Smith 1972) were used to measure the local infiltration capacities at sparse locations in the SP, before and after the flooding test was performed. A double-ring infiltrometer reproduces the local soil transition from unsaturated to saturated conditions that would take place under flooding conditions, and evaluates the infiltration capacity. In short, two metal rings are buried a few centimeters deep into the soil and filled with water (at constant or variable height). After a standard time of two hours, the infiltration rate approaches a steady state value, identified as $I_c(\mathbf{x}, t)$ at that location \mathbf{x} and time t.

During the flooding experiments it is possible to obtain infiltration values at the local scale by using other devices. For consistency we did not use other methods and rely on two double-ring campaigns, This implies that the only available measurements are those of $I_{c,0}$ and $I_{c,f}$. Only a few points are available for each one of these two variables. The former was obtained at 6 spatial points and the latter at those same 6 points plus three additional ones (a total of 9).

It is worthwhile noting that double-ring tests provide a direct estimate of infiltration capacity at the scale of the



Table 1 Experimental double ring test results performed in February 2009 (associated to $I_{c,f}$) and June 2009 (associated to $I_{c,f}$)

Location	$I_{c,0}$	$I_{c,f}$	Ratio
S1	0.19	0.18	0.94
S2	2.6	2.1	0.80
S3	2.9	2.5	0.86
S4	3.3	1.1	0.33
S5	12.9	1.2	0.09
S6	12.6	6.3	0.5
S7		0.17	
S8		3.04	
S9		0.75	
Mean	5.74	2.07	
Variance	30.6	3.98	

All values are expressed in m/d

device. In this case, they were representative of a small support scale ($\approx 0.12 \text{ m}^2$). This may have a significant influence in the geostatistical analysis, since it is important to use a method that accounts if necessary for the difference in support between the different variables involved. In Fig. 1 the "S" letters denote the locations where the double-ring infiltration tests were performed. The "C" letters refer to excavated pits where geology was directly observed and described.

Results from the analysis of the infiltration tests (using the modified Kostyakov method) are compiled in Table 1). It is clear that all points analyzed show a reduction in infiltration between the initial and the final stage. This reduction is larger (both in absolute and in relative terms) for the points displaying higher $I_{6.0}$ values.

3.3 Local scale modelling

The results of the correlation between the red band of the image presented in Fig. 1 and the logarithm of the infiltration values at $t = t_0$ and at $t = t_f$ are found to be linear (Fig. 3) according to Eqs. 4 and 5. A good linear correlation is obtained for both states of the artificial pond, with a Pearson's coefficient of $r^2 = 0.87$ for the February campaign $(t = t_0)$ and $r^2 = 0.89$ for the June campaign $(t = t_f)$.

3.4 Experimental variograms and cross-variograms

Field experiments are usually difficult to conduct and costly. Thus, one can typically always expect a very low number of data points, which in most cases (certainly in ours), prevents the direct estimation of the variogram of the infiltration capacity. Instead, a well behaving sampled variogram may be available for some secondary variable, P_{ν} in our case. Figure 4 shows the resulting two directional variograms obtained for P_{ν} in the principal directions. The direction of

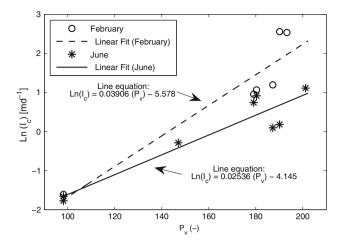


Fig. 3 Correlation between the pixel values (P_{ν}) from the red band of the satellite image (Fig. 1) and measured log-infiltration capacities (I_c) at selected locations at two different campaigns. The *lines* show the results from linear fitting to the experimental data. The Pearson's coefficient for the dotted line (February dataset) is $r^2 = 0.87$; for the straight line (June dataset), $r^2 = 0.89$

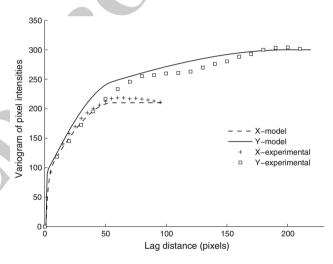


Fig. 4 Directional variograms of the pixel values in the reference image (exhaustive dataset)

maximum correlation is oriented along the y axis. The model variograms show three different structures, whose combined formulation in terms of principal directions is

$$\gamma_{P_{v}}(h_{x}) = 297 \left[0.32 \cdot \text{Gauss} \left(\frac{h_{x}}{7} \right) + 0.44 \cdot \text{Sph} \left(\frac{h_{x}}{92} \right) + 0.24 \cdot \text{Exp} \left(\frac{h_{x}}{400} \right) \right]$$

$$(10)$$

$$\gamma_{P_{v}}(h_{y}) = 297 \left[0.32 \cdot \text{Gauss} \left(\frac{h_{y}}{7} \right) + 0.44 \cdot \text{Sph} \left(\frac{h_{y}}{68} \right) \right]$$

$$(11)$$

where Gauss(.), Exp(.) and Sph(.) are the standard unitary variogram models (Deutsch and Journel 1998). 543





3.5 Mapping the clogging factor

Finally, the clogging factor λ should also be estimated. In our case, no direct field estimates about this parameter were measured. To overcome this problem, we assumed a perfect correlation between this parameter and the initial infiltration capacity Y_0 ,

$$Y_{\lambda} = a(Y_0) + b. \tag{12}$$

The coefficients a and b were estimated from the temporal evolution of the integrated infiltration capacity observed during the flooding test. The theory is unclear as to whether these two parameters should be positively or negatively correlated. Thus, we explore in our site the potential range of parameters in the power law formula (12) that can lead to a fit of the observed global infiltration behavior. In particular, we analyze the sign of the coefficient a. We will see how this may have a strong effect in devising potential remediation strategies.

Starting from given values a and b, the map of the λ values is obtained with (12), and then the local infiltration in space-time is obtained from (2). Finally, the integrated infiltration is obtained with (3). The method is repeated by changing the values of a and b until the curve matches that of Fig. 2.

Several potential combinations of the two parameters might lead to reasonable fits. In Fig. 5 we fix two different b values and explore the sensitivity of the global infiltration curves to a. It is found that we are able to obtain two sets of parameters (a, b), (-0.4, -0.3) and (0.3, -3.0) that lead to a similar fit. It can be observed that the curves are quite sensitive to the two parameters, providing a way to calibrate them. Calibration cannot be fully completed in our site since no data for intermediate times could be recorded. As a consequence we believe that there is a need to perform short flooding tests in real sites in order to obtain real values of λ that could be used to design an optimal management operation. This can only be obtained by performing test during the flooding period. The problem is that infiltration measurements are quite sensitive to the method used, and so it is difficult to be able to combine data coming from double-ring or seepage meters, for example.

While one can find different sets of (a, b) values, the key is the sign of the a parameter. Depending on this sign we will have enhanced clogging in regions of initially high infiltration capacities, or alternatively, clogging enhanced in points displaying initially small infiltration capacities. This can be seen in Fig. 6, which includes the maps corresponding to the two sets of (a, b) already presented before. By construction, the maps are visually highly correlated with those of initial or final infiltration, included also in the same figure.

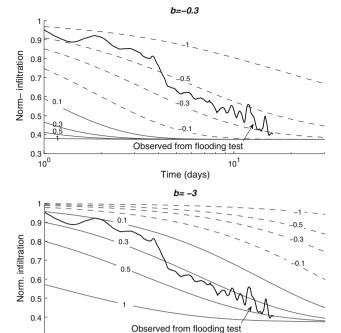


Fig. 5 Sensitivity of the correlation model between $\ln(I_c)$ and $\ln(I_0)$ on the slope coefficient a and the y-intercept. The results are plotted for b=-0.3 (up) and b=-3 (down). The normalized observed values, calculated using the water balance equation, are also plotted for comparison with the curves

Time (days)

0.3

One of the main factors controlled by a is the actual shape of the pdf of λ . Thus, while the two sets of parameters lead to a very similar fit of the global infiltration curve, t actual values of λ are very different (see Fig. 7). A direct implication is that it is not possible to derive an equivalent λ value from the local ones capable of reproducing the full behavior. This can be observed in Fig. 8, where the evolution of infiltration with time using the harmonic, geometric and arithmetic means of the point values presented in Fig. 7 are computed. It is clear that upscaling from local values would be an error in general, and that some conditioning on real values is needed to produce a proper reproduction of the global behavior.

The evolution of the pdf of infiltration capacity with time is also dependent on the map of λ values. This can be observed in Fig. 9, where a positive a value implies a slower reduction in the infiltration values with time than a negative a value. This is caused by the former leading to very small λ values in Fig. 7. These small values cause infiltration to be reducing slowly with time. The picture completely changes for the alternative set of parameters (-0.4, -0.3) where the local λ values are high, and thus infiltration reduces very fast with time.

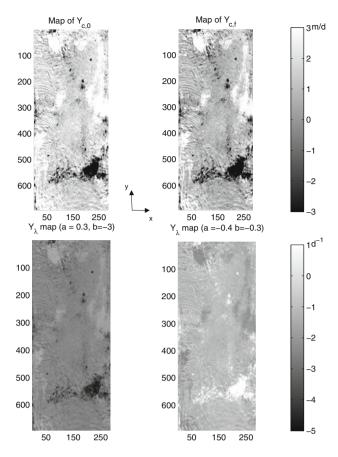


Fig. 6 Cokriged maps of the logarithm of the infiltration capacities of the pilot SP at t_0 and $t_f(up)$, and corresponding λ maps using positive a on the *left* and negative a on the *right*

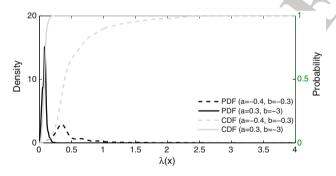


Fig. 7 Probability density function (pdf, black lines) and cumulative frequency (cdf, gray lines) of $\lambda(\mathbf{x})$ calculated using different combination of a and b values, and mapped in Fig. 4 (down)

These results will be significant when devising a remediation method. Notice how in Fig. 9 it is observed that for the set (0.3, -3.0) all values of infiltration decrease more or less equally in time. Thus, the best alternative for remediation would be to empty the full pond and restore the initial capacity if possible. On the other hand, for the (-0.4, -0.3) we observe how the reduction of infiltration is

mostly produced by the fast reduction in the more permeable areas. Thus, we could devise a method that targets the initially high infiltration areas. If clogging in these points is prevented, the global infiltration will remain high for a long period of time.

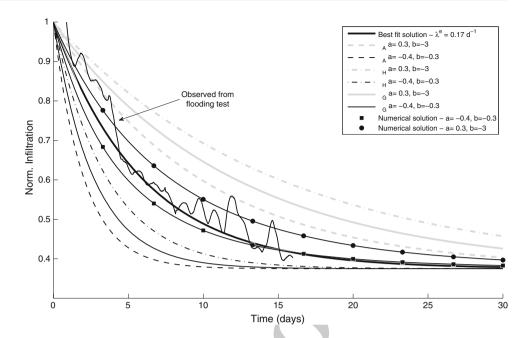
4 Conclusions

Assessing spatio-temporal variability of soil hydraulic variables is often needed to make optimal, effective and efficient decisions in many branches of the groundwater hydrology. In systems with complex processes and in a heterogeneous geological context, such as in the case of artificial recharge practices using surface ponds (SP), the estimation of such parameters is highly uncertain. This is mainly due to scarce primary information, due to the high costs of direct field observation, due to limited accessibility to the sites or even due to partial knowledge of the physical processes involved. Lumped models are often preferable to work with due to their simplicity and versatility, and secondary information could provide reliable supporting data to be used for mapping purposes of primary variables.

In this paper, we describe a methodology to map the spatio-temporal distribution of the characteristic lumped temporal factors of clogging (λ) in a SP, with the aid of satellite images and a few experimental datasets based on local sparse measurements. This allows us to estimate the spatio-temporal variability of the infiltration capacities of the topsoil. This potentially allows managers to find optimal strategies for clean-up operations that can minimize the effect of clogging without having to operate the full facility and to make educated decisions that might mitigate the failure in O&M operations in an SP.

As a case study, we applied this method to a pilot site in Spain. Although the developed model introduces some simple hypotheses and simplifications in this illustrative example, we showed that we were able to map the evolution of local infiltration with time. To validate the method, we compared the measured total infiltration capacities at the SP scale with the ones calculated by integration of calculated local values using a geostatistical-physically-mixed approach. Numerical results satisfactorily agree with the observations, accounting for the multiple necessary approximations we assumed.

It is found that arguably the most significant parameter is the sign of the a parameter in the power law model relating λ and the initial infiltration. Depending on this sign it is possible to see the practical relevance of designing a remediation method that targets only parts of the domain (acting on the high infiltrating areas during operation), or else it is better to use the classical approach of treatment



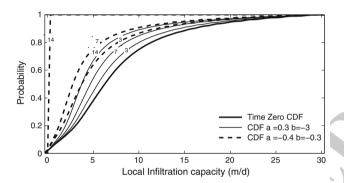


Fig. 9 Evolution of the probability of local infiltration capacity values in the pond at different times (the tags express the days elapsed after the flooding test began) using two possible combination of relationships between λ and the initial infiltration capacity ____

after emptying the facility. This alternative might have significant impact in managing an SP site.

Acknowledgments This project was financially by the Spanish Ministry of Science and Innovation via the CONSOLIDER-Ingenio 2010 (CSD2009-00065), RARA-AVIS (CGL2009-11114) and HEROS (CGL2007-66748) projects. DP would like to further acknowledge the Spanish government for funding through the FPU scholarship programme (Formation of University Lecturers).

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