

MATERHORN

Summary of Progress at NPS and NCAR*

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*Does not represent work by Jason Knievel and Yubao Liu in collaboration with UU and UVA

Topics



- Data assimilation (ensemble)
- Ensemble sensitivities for observing strategies)
- Model error estimation
- Parameter estimation
- Observation bias estimation



Ensemble Sensitivities

Network design for fine-scale near-surface
forecasts in complex terrain

Ensemble sensitivity analysis (ESA)

How does the change in a set of initial state variables x_s change a forecast metric J ?

$$\frac{\partial J_e}{\partial x^a}$$

- Identify dynamically relevant covariance structures in space and time
- **Propose observing strategies for mesoscale, short-range forecasts in complex terrain**
- Sensitivity scales (time and space) to infer predictability scales
- Predictability of specific phenomena
- Open issues:
 - Sampling error
 - Linearity assumptions in complex terrain

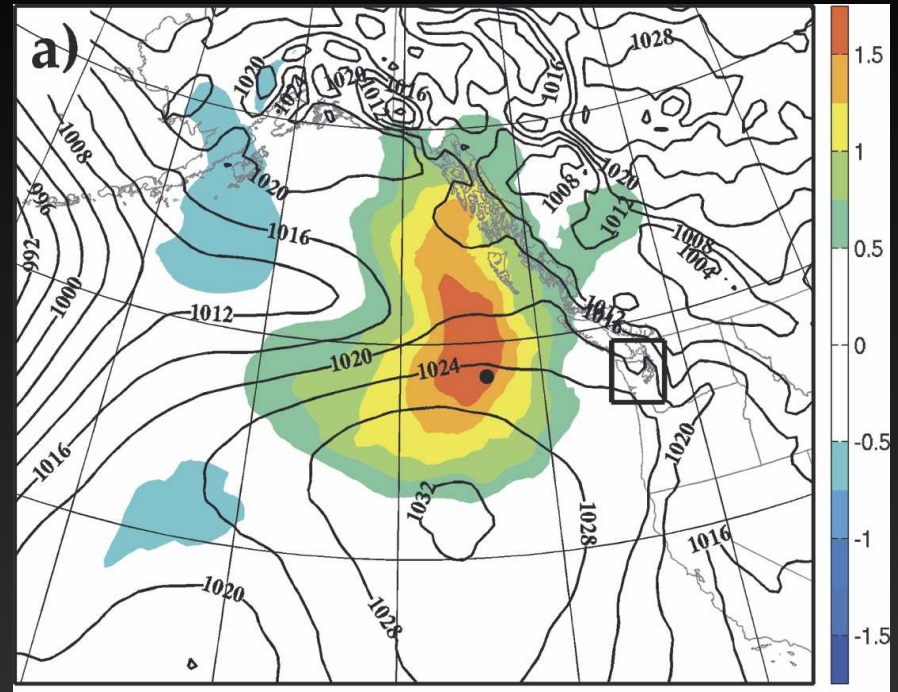
Ensemble Sensitivity Background



NCAR

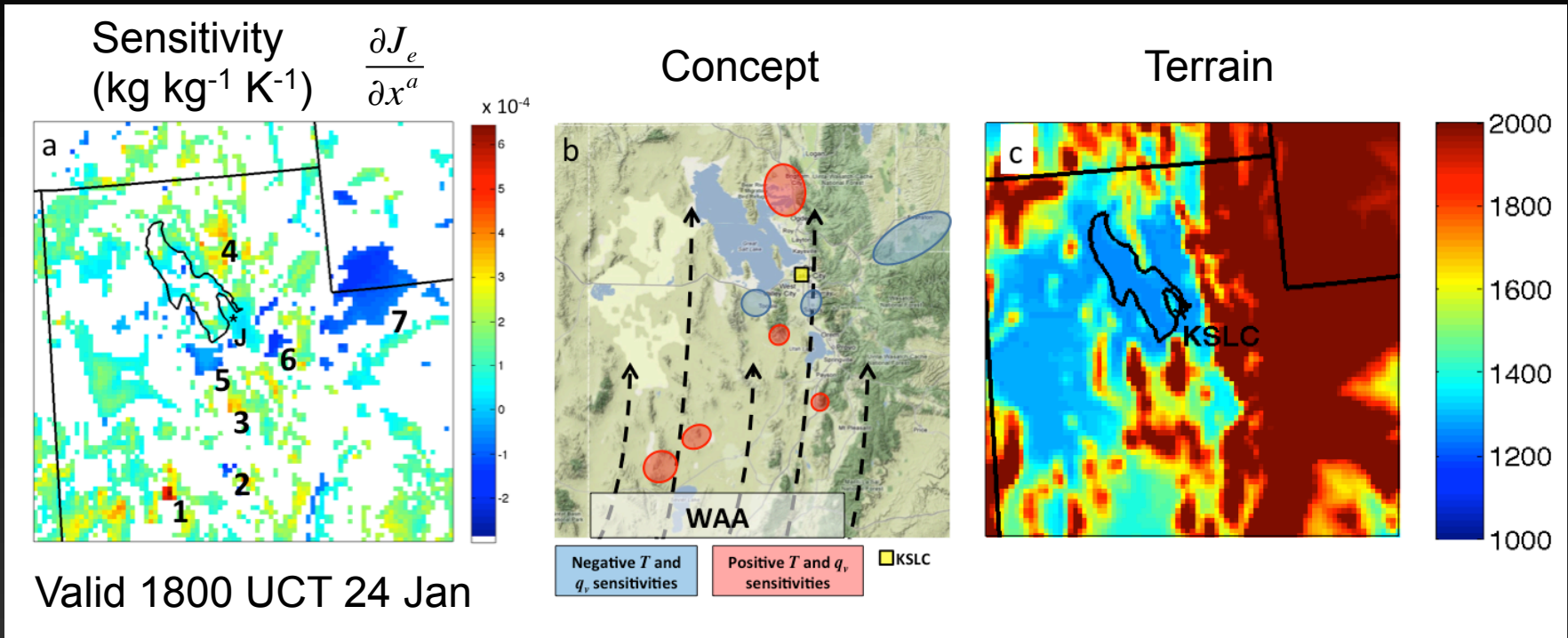
- Ancell and Hakim (2007) showed theoretical equivalence between adjoint and ensemble sensitivity for linear perturbations and Gaussian statistics
- Relies on linearization about an ensemble-mean trajectory
- Rigorous application has so far been limited to large-scale (smooth) and integrated processes where strong linear relationships are more likely

An optimal ensemble data assimilation system provides an appropriate sample



Sensitivity of 24-h sea-level pressure (SLP) over western Washington to SLP initial conditions, and ensemble-mean SLP (from Torn and Hakim 2008).

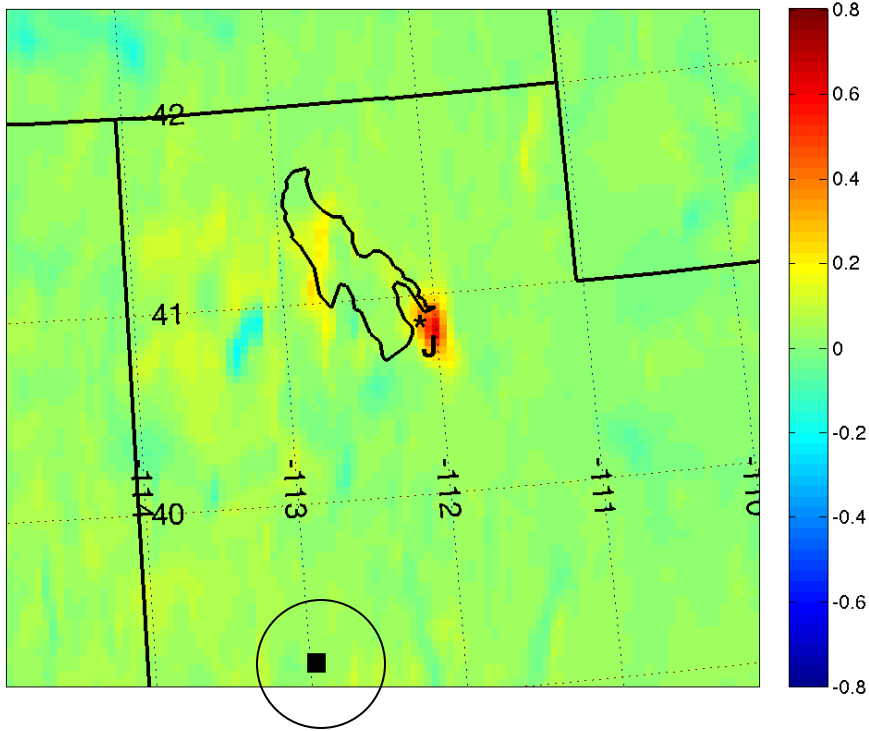
Moisture sensitivity to temperature



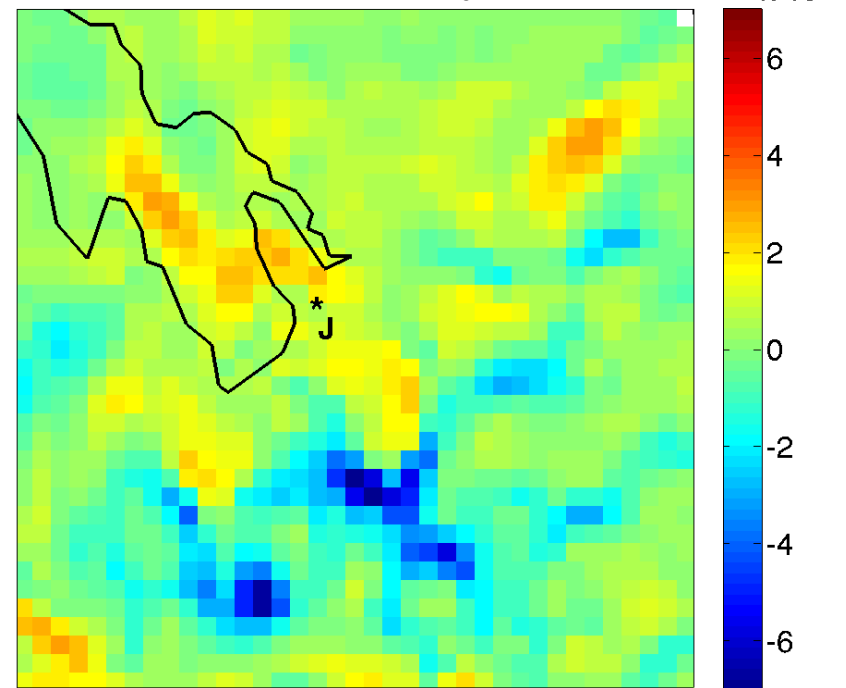
$J = 2 \times 2 \times 2$ box-mean water vapor mixing ratio over Salt Lake City airport
 $x =$ Potential temperature (here on model first layer)

Perturbation experiments

Analysis perturbation, $\theta(K)$



Forecast perturbation t_0+6h ($kg\ kg^{-1}$)



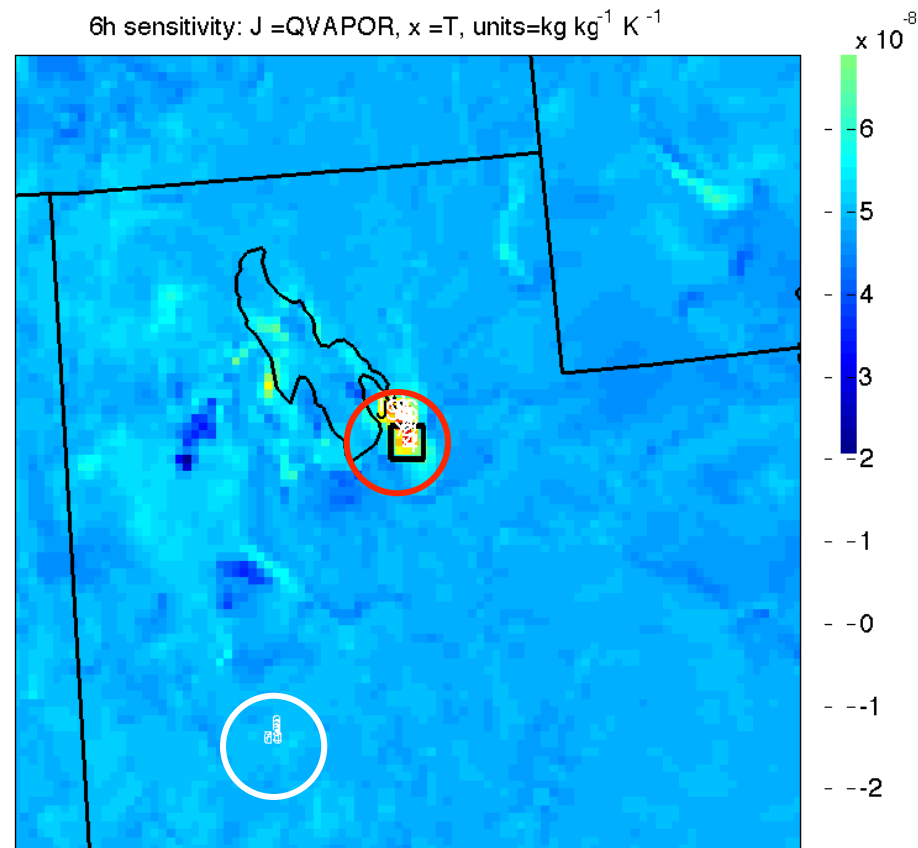
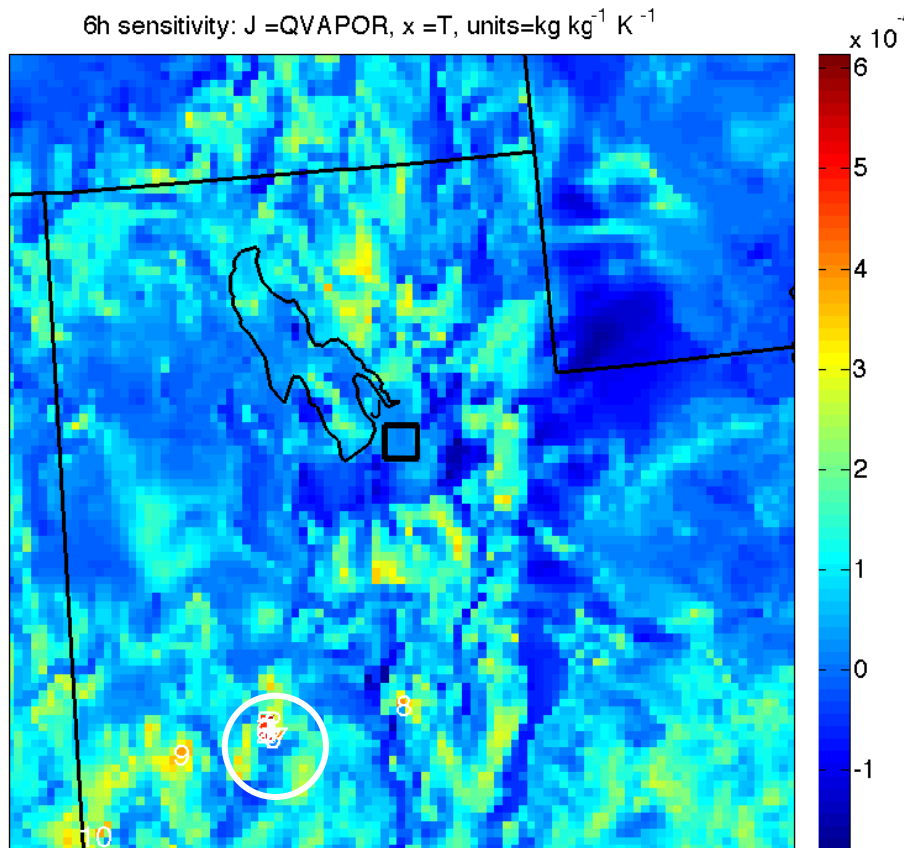
Perturbation of one analysis standard deviation in θ at the most sensitive location, *assimilated* with ensemble filter.

Effect of approximation

Diagonal approximation

Full covariance

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Approximation under-emphasizes sensitivities local to the response. Agreement on some sensitive points (numbered) to southwest of response.

Ensemble Sensitivity with Localization

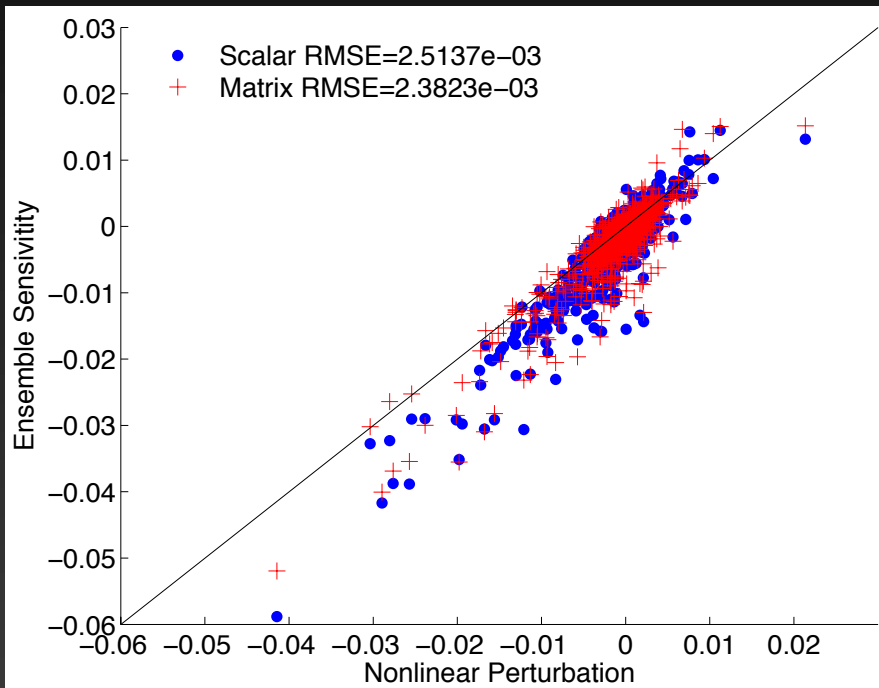
$$\begin{aligned}\delta J &= \alpha \circ \left\{ \mathbf{J}_e^T \left[\mathbf{X}_i^a \left(\mathbf{X}_i^{aT} \mathbf{X}_i^a \right)^{-1} \right]^T \rho \circ \mathbf{P}_i^a \mathbf{h}_{i+1}^T \left(\mathbf{h}_{i+1} \rho \circ \mathbf{P}_i^a \mathbf{h}_{i+1}^T + \mathbf{R} \right)^{-1} \left(y_{i+1}^o - \mathbf{h}_{i+1} \mathbf{X}_i^a \right) \right\} \\ &= \alpha \circ \left\{ \mathbf{J}_e^T \left[\mathbf{X}_i^a \left(\mathbf{X}_i^{aT} \mathbf{X}_i^a \right)^{-1} \right]^T \delta \mathbf{X}^a \right\}\end{aligned}$$

- Covariance localization, or tapering, can be applied
 - at the assimilation step with ρ
 - to the regressions with α
- ρ is typically a function of space alone
- α is function of space and time, here from a Bayesian hierarchical estimate (Anderson 2007)

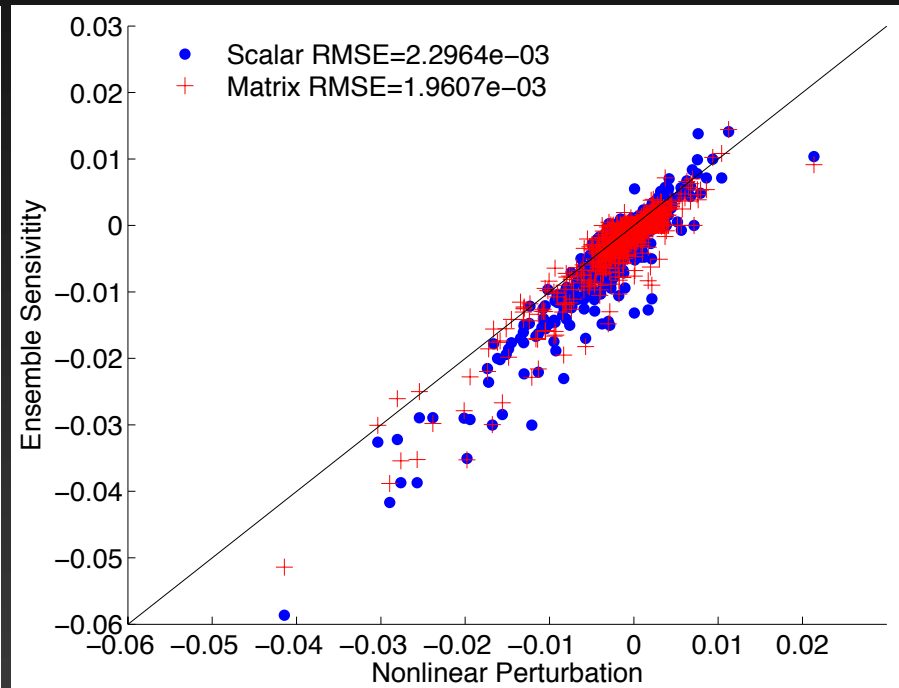
Perfect Model (two scales)

When only smooth/slow scales present, little difference between univariate (scalar) and multivariate (matrix) predictions of response to perturbation.

Sensitivity *without* localization



Sensitivity *with* localization



Here observations are randomly chosen from every other gridpoint (which are un-observed for sensitivity calculations).

Imperfect Model

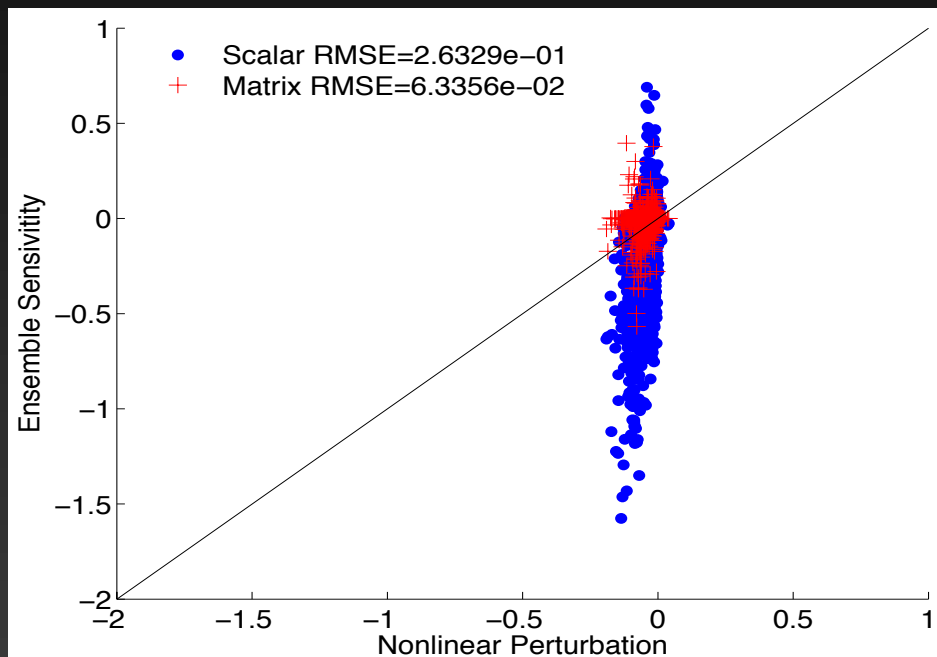
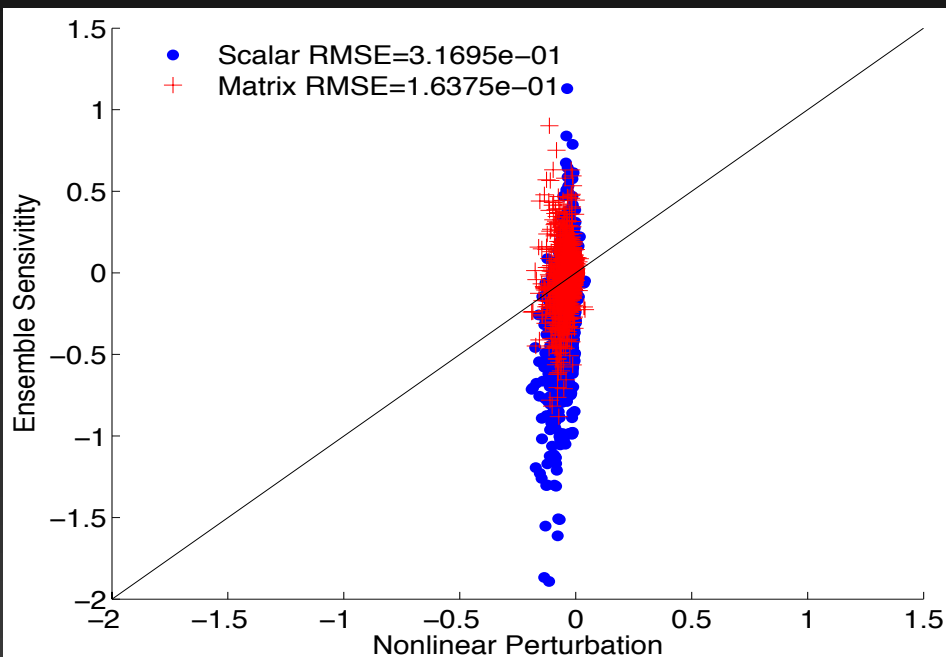


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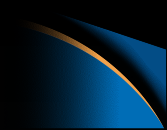
For imperfect model, diagonal approximation results in greater over-prediction of response; multivariate sensitivities account for presence of fast scales in real system, which appears as noise.

Sensitivity *without* localization

Sensitivity *with* localization



Here observations are assimilated on half of domain that is data void; more impact from observations because greater uncertainty in analysis.



Model Inadequacy

Quantifying model error with ensemble data
assimilation

Bias in state estimation



A scalar example (adapted from Dee and DaSilva 1998):

Consider $\beta_f = E(\varepsilon_f) \neq 0$
 $\beta_o = E(\varepsilon_o) = 0$ The background/forecast is biased but the observation is not.

Then $\sigma_f^2 = E[(\varepsilon_f - \beta_f)^2]$
 $\sigma_o^2 = E(\varepsilon_o^2)$ The background error variance is biased, but the observation error variance is not.

If β_b is considered, the analysis is **unbiased**:

$$\text{for } \sigma_f^2 = \sigma_o^2 = \sigma^2 : \\ x_a = \frac{1}{2}(x_f - \beta_f + y_o) \\ \beta_a = 0, \quad E(\varepsilon_a^2) = \frac{1}{2}\sigma^2$$

If β_b is ignored, the analysis is **biased**:

$$\text{for } \sigma_f^2 = \sigma_o^2 = \sigma^2 : \\ x_a = \frac{1}{2}(x_f + y_o) \\ \beta_a = \frac{1}{2}\beta_f, \quad E(\varepsilon_a^2) = \frac{1}{4}\beta_f^2 + \frac{1}{2}\sigma^2$$

Systematic model errors

- Persistent structures or time-means in the innovations or increments in data assimilation, or tendencies in the model, result from systematic error.
- Objective data assimilation cannot eliminate bias.

For $\sigma_f^2 = \sigma_o^2 = \sigma^2$ and an unbiased observation:

$$x_a - x_f = \frac{1}{2}(x_f + y_o) - x_f = \frac{1}{2}(y_o - x_f)$$

$$\begin{aligned} 2E(x_a - x_f) &= E(y_o - x_f) = E(y_o - x_t - \varepsilon_f) = [E(y_o) - E(x_t) - \beta_f] \\ &= [E(x_t) + E(\varepsilon_o) - E(x_t) - \beta_f] = \beta_f \end{aligned}$$

- Bias can β be estimated from analysis increments or background.
- Not immediately clear whether the observations or the model is biased, but we can extract the statistics given an observation network containing unbiased observations.

Systematic observation errors

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For $\sigma_f^2 = \sigma_o^2 = \sigma^2$ and an unbiased forecast:

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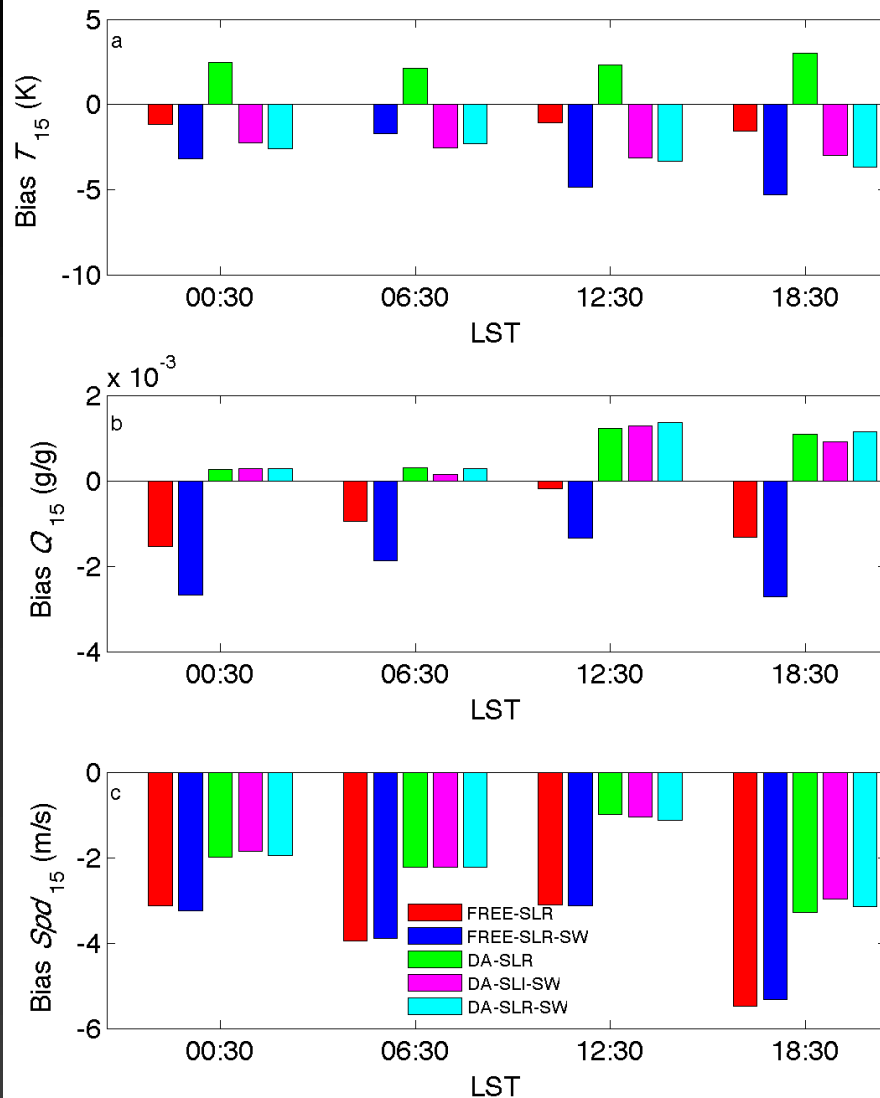
$$\begin{aligned} 2E(x_a - x_f) &= E(y_o - x_f) = E(y_t - \varepsilon_f - x_f) = [E(y_t) + \beta_o - E(x_f)] \\ &= [E(x_t) + \beta_o - E(x_t) - E(\varepsilon_f)] = \beta_o \end{aligned}$$

- Bias can β be estimated from analysis increments or background.
- Not immediately clear whether the observations or the model is biased, but we can extract the statistics given an observation network containing unbiased observations.

Predicting (independent) observations at layer-1 (~15 m AGL)



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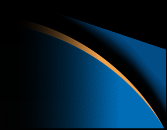
- Bias magnitudes of winds smaller when assimilating
- Biases here are generally the same sign as biases in observation space

Nothing looks too strange

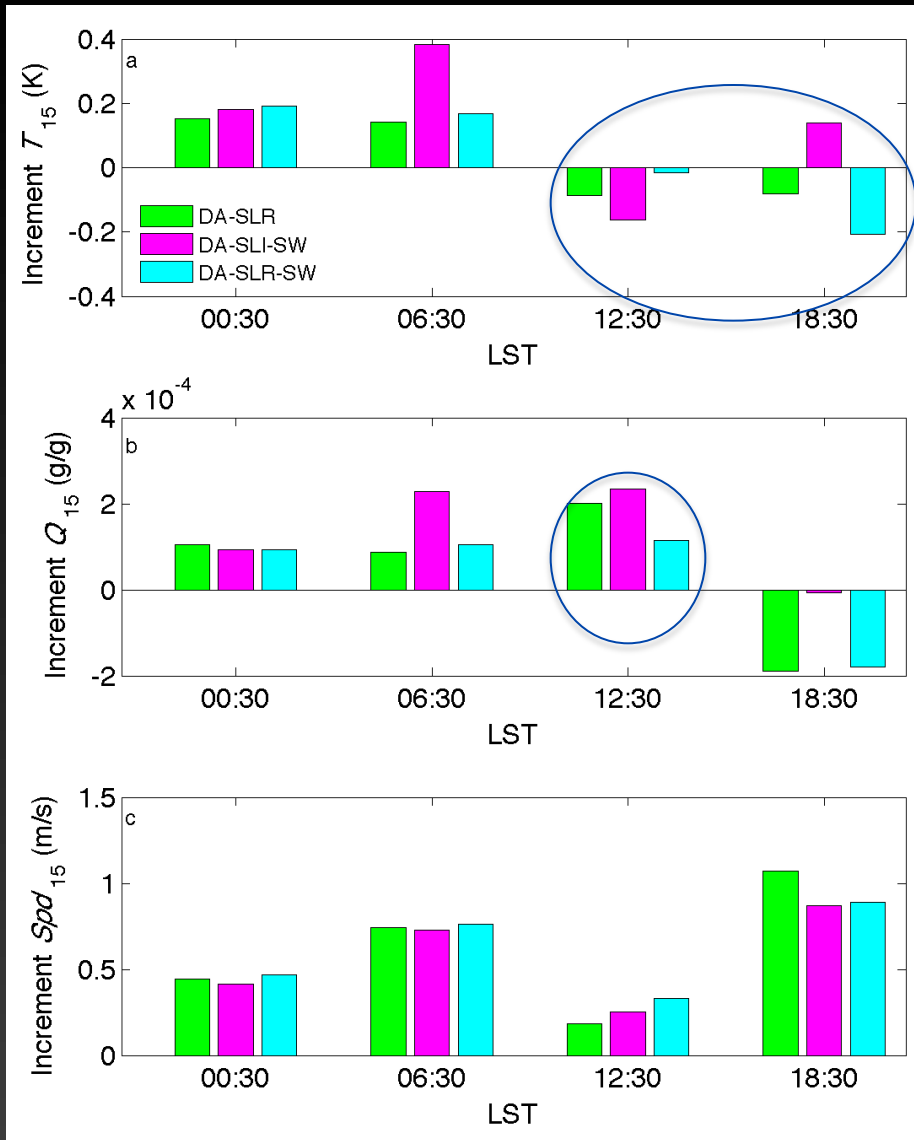
- Bias generally of the same sign as at shelter or anemometer
- Magnitude of bias not reduced as clearly as at shelter or anemometer
- Except: FREE-SLR less biased in temperature than DA-SLR (maybe a little strange)

← Night Day →

Analysis increments



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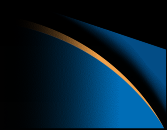
Signs that covariances can be biased:

- Some increments are acting to increase the bias

Possible causes:

- Nonlinear statistics in regressions underlying the assimilation (not the case here)
- Poor parameterization of surface-layer profiles

← Night Day →



Model Inadequacy

Parameter estimation to address parametric model error (and systematic observation error)

Parametric errors



Theory leads to state augmentation approach:

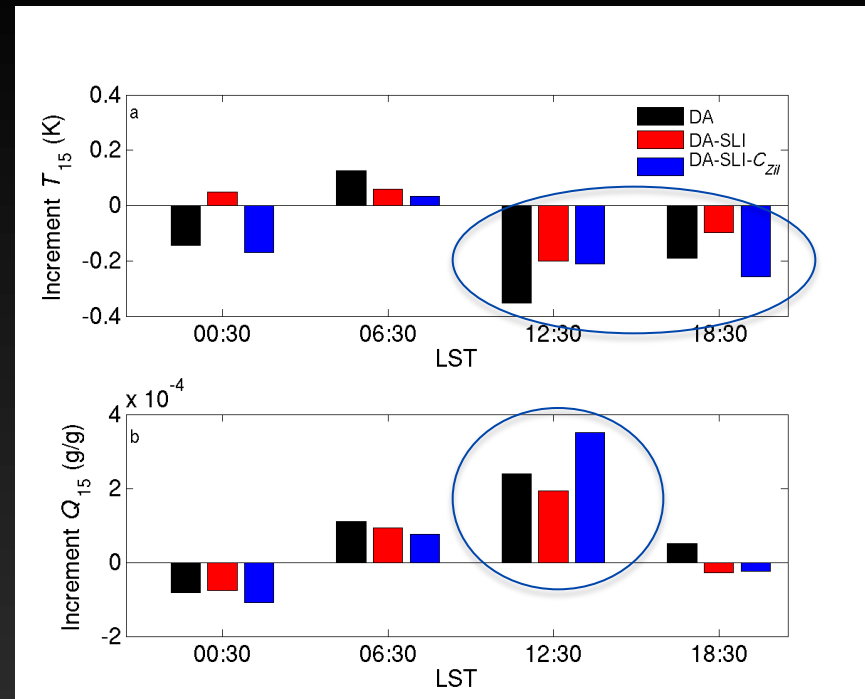
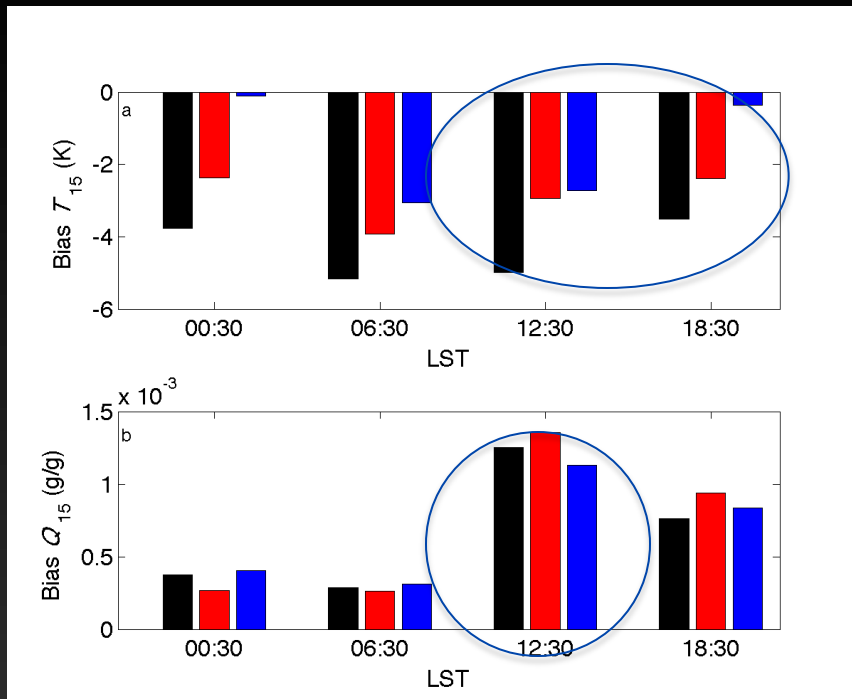
$$\mathbf{z} = \begin{bmatrix} \mathbf{x} \\ \mathbf{p} \end{bmatrix}$$

“Augmented” state vector
State vector
 \mathbf{x} Vector of parameters
 \mathbf{p}

The statistical analysis equation is still valid; forward operators and covariance matrices are modified accordingly

- If a model prediction is sensitive to a parameter value, then covariances between parameters and predicted observations (\mathbf{Gz}_b) will result (\mathbf{G} is the augmented forward operator).
- Parameter values can be modified by observations.
- **Parameters can be in the model or otherwise arbitrary.**

Bias and analysis increments

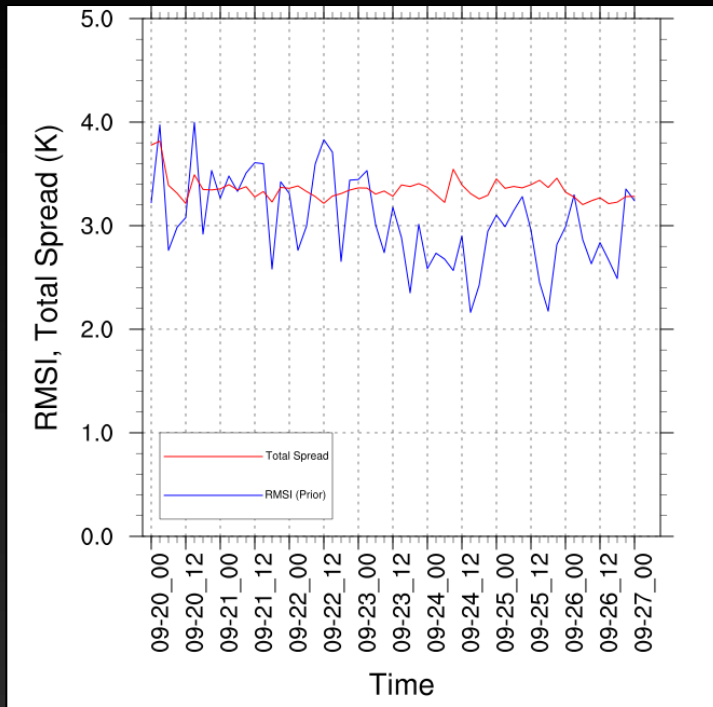


- Even with parameter estimation, analysis increments can be systematically in the wrong direction.
- DA increasing the biases via the analysis increments.
- Suggests structural errors.

Parameter estimation (3D model)



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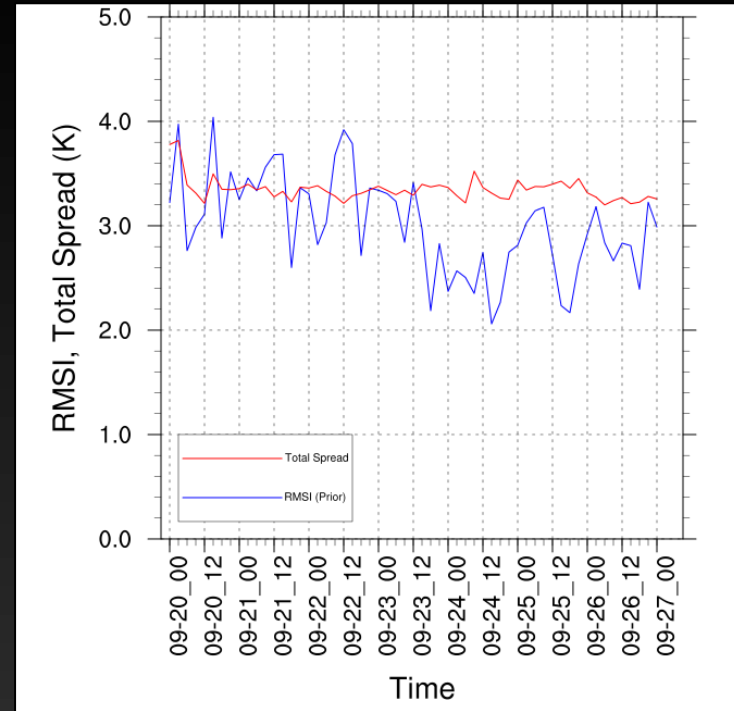
METAR 2-m T

Constant C_{ZIL}

Mean RMS Innov. = 3.01 K

Mean Tot. Spread = 3.33 K

(Means discard 20 Sep)



METAR 2-m T

Variable C_{ZIL}

Mean RMS Innov. = 2.96 K

Mean Tot. Spread = 3.33 K

(Means discard 20 Sep)

Forward operators



- Give predicted observations from forecasts model prognostic variables: $\mathbf{H}\mathbf{x}^f \approx h(\mathbf{x}^f)$
- Simple:
 - Interpolation to a radiosonde observation
 - Diagnostic surface-layer variables (2-m T, 10-m winds) and interpolation to horizontal location)
- Complex:
 - Satellite radiances (also nonlinear)
 - GPS radio occultation (also nonlinear)
 - Doppler winds
 - Radar reflectivity (also nonlinear)

Observing scales

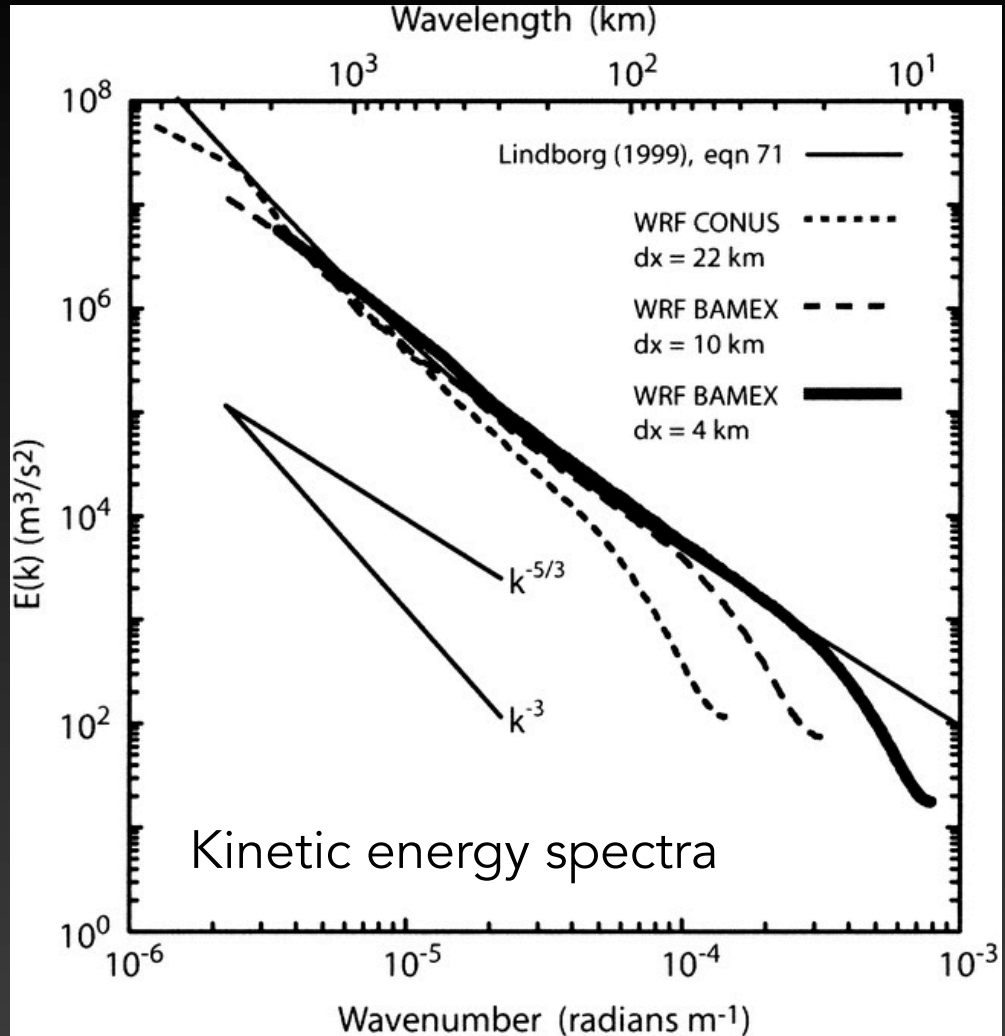


- An observation “sees” all scales of motion slower than its sampling rate.
- Most observations sample well below the numerically diffusive range.
- Most useful observations contain energy projecting onto scales above the diffusive range.

Examples:

- 10-minute average surface obs
- Packages on small UAVs.

Skamarock (2004)



Progress summary



- Ensemble sensitivities for observation network design in complex terrain
- Model error quantification and parameter estimation for land-atmosphere coupling errors
- Parameter estimation for near-surface observation error estimation (Raquel Lorente-Plazas)

References



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- **Hacker, J. P.** and L. Lei, 2015: Multivariate ensemble sensitivity with localization. *Mon. Wea. Rev.*, **143**, 2013–2027. doi: <http://dx.doi.org/10.1175/MWR-D-14-00309.1>
- Wile, S.M., **J. P. Hacker**, and K.H. Chilcoat, 2015: The potential utility of high-resolution ensemble sensitivities during weak flow in complex terrain. Accepted for publication in *Wea. and Forecast.*, May 2015.
- Ryerson, W., and **J. Hacker**, 2014: The potential for mesoscale visibility predictions with a multi-model ensemble. *Wea. and Forecast.*, **29**, 543–562. doi: <http://dx.doi.org/10.1175/WAF-D-13-00067.1>
- **Hacker, J.** and W. Angevine, 2013: Ensemble data assimilation to characterize surface-layer errors in numerical weather prediction models. *Mon. Wea. Rev.*, **141**, 1804–1821. doi: <http://dx.doi.org/10.1175/MWR-D-12-00280.1>

Statistical-analysis equation



$$\begin{aligned}\mathbf{x}_a &= \mathbf{x}_b + \mathbf{K}(\mathbf{y}_o - \mathbf{H}\mathbf{x}_b) \\ \mathbf{K} &= \frac{\mathbf{B}\mathbf{H}^T}{\mathbf{H}\mathbf{B}\mathbf{H}^T + \mathbf{R}} \\ \mathbf{B} &= \langle \mathbf{e}_b \mathbf{e}_b^T \rangle, \quad \mathbf{R} = \langle \mathbf{e}_o \mathbf{e}_o^T \rangle \\ \mathbf{P}_a &= \langle \mathbf{e}_a \mathbf{e}_a^T \rangle = (\mathbf{I} - \mathbf{K}\mathbf{H})\mathbf{B}\end{aligned}$$

Under assumptions of Gaussian, unbiased errors, and linear operator \mathbf{H} , this equation holds for:

- Optimal interpolation
- Variational minimization
- Kalman filters
- Other linear filters

Differences arise when trying to estimate the error statistics, or when additional constraints are imposed.

Effect of hypothetical θ observation

$$\delta J_e = \frac{\partial J_e}{\partial x^a} \mathbf{K} (\mathbf{y}^o - \mathbf{h} \mathbf{x}^a)$$
$$\mathbf{K} = \mathbf{P}^a \mathbf{h}^T (\mathbf{h} \mathbf{P}^a \mathbf{h}^T + \mathbf{R})^{-1}$$

Can test use of sensitivities to predict the change in forecast metric resulting from a hypothetical observation.

Analysis increment can come from:

- assimilating synthetic obs
- approximation with univariate linear regression

