

“Moderated Mediation Analysis Using Bayesian Methods”

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Appendix A

Testing whether a mediation effect is moderated using the delta method

This appendix contains hypotheses and derivations of point estimation ($f(\hat{\theta})$) and standard error ($s.e.(f(\hat{\theta}))$) for testing whether a mediation effect is moderated for the 6 scenarios presented in Table 1. The first-order and second-order variance estimates of the point estimate based on the delta method can be derived by $\mathbf{D}'\hat{\Sigma}(\hat{\theta})\mathbf{D}$ and $\mathbf{D}'\hat{\Sigma}(\hat{\theta})\mathbf{D} + \frac{1}{2}tr\{(\mathbf{H}\hat{\Sigma}(\hat{\theta}))^2\}$ (see Preacher et al., 2007), where $\mathbf{D} = \partial_{\theta}(f(\hat{\theta}))$ and $\mathbf{H} = \partial_{\theta}^2(f(\hat{\theta}))$. Therefore, the first-order test statistic is

$$Z = \frac{f(\hat{\theta})}{\sqrt{\mathbf{D}'\hat{\Sigma}(\hat{\theta})\mathbf{D}}} \text{ and the second-order test statistic is } Z = \frac{f(\hat{\theta})}{\sqrt{\mathbf{D}'\hat{\Sigma}(\hat{\theta})\mathbf{D} + \frac{1}{2}tr\{(\mathbf{H}\hat{\Sigma}(\hat{\theta}))^2\}}}.$$

Scenario 0: no moderation

The null hypothesis is $H_0 : ab = 0$ with a point estimate of $f(\hat{\theta}) = \hat{a}\hat{b}$. The first-order and second-order variance estimates of the point estimate are $\hat{b}^2 s_a^2 + \hat{a}^2 s_b^2$ and $\hat{b}^2 s_a^2 + \hat{a}^2 s_b^2 + s_a^2 s_b^2$.

Therefore, the first-order test statistic is $Z = \frac{\hat{a}\hat{b}}{\sqrt{\hat{b}^2 s_a^2 + \hat{a}^2 s_b^2}}$ and the second-order test statistic is

$$Z = \frac{\hat{a}\hat{b}}{\sqrt{\hat{b}^2 s_a^2 + \hat{a}^2 s_b^2 + s_a^2 s_b^2}}.$$

Scenario 1: b is moderated by the input variable X

The null hypothesis is $H_0 : ab_2(x_1 - x_2) = 0$ with a point estimate of $f(\hat{\theta}) = \hat{a}\hat{b}_2(x_1 - x_2)$.

The first-order and second-order variance estimates of the point estimate are

$(\hat{b}_2^2 s_a^2 + \hat{a}^2 s_{b_2}^2)(x_1 - x_2)^2$ and $(\hat{b}_2^2 s_a^2 + \hat{a}^2 s_{b_2}^2 + s_a^2 s_{b_2}^2)(x_1 - x_2)^2$. Therefore, the first-order test statistic is $Z = \frac{\hat{a}\hat{b}_2}{\sqrt{\hat{b}_2^2 s_a^2 + \hat{a}^2 s_{b_2}^2}}$ and the second-order test statistic is $Z = \frac{\hat{a}\hat{b}_2}{\sqrt{\hat{b}_2^2 s_a^2 + \hat{a}^2 s_{b_2}^2 + s_a^2 s_{b_2}^2}}$.

Scenario 2: a is moderated by a moderator variable Z but b is not moderated by Z

The null hypothesis is $H_0 : a_2 b(z_1 - z_2) = 0$ where z_1 and z_2 are two values of the moderator variable Z . The point estimate is $\hat{a}_2 \hat{b}(z_1 - z_2)$ and the first-order and second-order variance estimates of the point estimate are $(\hat{a}_2^2 s_b^2 + \hat{b}^2 s_{a_2}^2)(z_1 - z_2)^2$ and

$(\hat{a}_2^2 s_b^2 + \hat{b}^2 s_{\hat{a}_2}^2 + s_{\hat{a}_2}^2 s_b^2)(z_1 - z_2)^2$. Therefore, the first-order test statistic is $Z = \frac{\hat{a}_2 \hat{b}}{\sqrt{\hat{a}_2^2 s_b^2 + \hat{b}^2 s_{\hat{a}_2}^2}}$ and the second-order test statistic is $Z = \frac{\hat{a}_2 \hat{b}}{\sqrt{\hat{a}_2^2 s_b^2 + \hat{b}^2 s_{\hat{a}_2}^2 + s_{\hat{a}_2}^2 s_b^2}}$.

Scenario 3: b is moderated by Z but a is not moderated by Z

The null hypothesis is $H_0 : ab_2(z_1 - z_2) = 0$ where z_1 and z_2 are two values of the moderator variable Z . The point estimate is $\hat{a}\hat{b}_2(z_1 - z_2)$ and the first-order and second-order variance estimates of the point estimate are $(\hat{a}^2 s_{\hat{b}_2}^2 + \hat{b}_2^2 s_a^2)(z_1 - z_2)^2$ and $(\hat{a}^2 s_{\hat{b}_2}^2 + \hat{b}_2^2 s_a^2 + s_a^2 s_{\hat{b}_2}^2)(z_1 - z_2)^2$. Therefore, the first-order test statistic is $Z = \frac{\hat{a}\hat{b}_2}{\sqrt{\hat{a}^2 s_{\hat{b}_2}^2 + \hat{b}_2^2 s_a^2}}$ and the second-order test statistic is $Z = \frac{\hat{a}\hat{b}_2}{\sqrt{\hat{a}^2 s_{\hat{b}_2}^2 + \hat{b}_2^2 s_a^2 + s_a^2 s_{\hat{b}_2}^2}}$.

Scenario 4: a is moderated by Z_1 and b is moderated by Z_2

The null hypothesis is $H_0 : a_1 b_2(z_{21} - z_{22}) + a_2 b_1(z_{11} - z_{12}) + a_2 b_2(z_{11} z_{21} - z_{12} z_{22}) = 0$ where z_{11} and z_{12} are two values of the moderator variable Z_1 and z_{21} and z_{22} are two corresponding values of the moderator variable Z_2 . Here, we have

$$\hat{\boldsymbol{\theta}} = [\hat{a}_1 \hat{a}_2 \hat{b}_1 \hat{b}_2]$$

$$f(\hat{\boldsymbol{\theta}}) = \hat{a}_1 \hat{b}_2(z_{21} - z_{22}) + \hat{a}_2 \hat{b}_1(z_{11} - z_{12}) + \hat{a}_2 \hat{b}_2(z_{11} z_{21} - z_{12} z_{22})$$

$$\hat{\boldsymbol{\Sigma}}(\hat{\boldsymbol{\theta}}) = \begin{bmatrix} s_{\hat{a}_1}^2 & s_{\hat{a}_1 \hat{a}_2} & 0 & 0 \\ s_{\hat{a}_1 \hat{a}_2} & s_{\hat{a}_2}^2 & 0 & 0 \\ 0 & 0 & s_{\hat{b}_1}^2 & s_{\hat{b}_1 \hat{b}_2} \\ 0 & 0 & s_{\hat{b}_1 \hat{b}_2} & s_{\hat{b}_2}^2 \end{bmatrix}$$

and thus

$$\begin{aligned} \mathbf{D}'\hat{\boldsymbol{\Sigma}}(\hat{\boldsymbol{\theta}})\mathbf{D} &= s_{\hat{b}_1}^2 \hat{a}_2^2 (z_{11} - z_{12})^2 + s_{\hat{a}_1}^2 \hat{b}_2^2 (z_{21} - z_{22})^2 + 2(\hat{a}_1 \hat{a}_2 s_{\hat{b}_1 \hat{b}_2} + \hat{b}_1 \hat{b}_2 s_{\hat{a}_1 \hat{a}_2})(z_{11} - z_{12})(z_{21} - z_{22}) \\ &+ 2\hat{a}_2^2 s_{\hat{b}_1 \hat{b}_2} (z_{11} - z_{12})(z_{11} z_{21} - z_{12} z_{22}) + 2\hat{b}_2^2 s_{\hat{a}_1 \hat{a}_2} (z_{21} - z_{22})(z_{11} z_{21} - z_{12} z_{22}) \\ &+ s_{\hat{b}_2}^2 (\hat{a}_1 (z_{21} - z_{22}) + \hat{a}_2 (z_{11} z_{21} - z_{12} z_{22}))^2 + s_{\hat{a}_2}^2 (\hat{b}_1 (z_{11} - z_{12}) + \hat{b}_2 (z_{11} z_{21} - z_{12} z_{22}))^2 \end{aligned}$$

$$\begin{aligned} \frac{1}{2}tr\{(\mathbf{H}\hat{\Sigma}(\hat{\boldsymbol{\theta}}))^2\} &= s_{\hat{a}_1}^2 s_{\hat{b}_2}^2 (z_{21} - z_{22})^2 + s_{\hat{a}_2}^2 s_{\hat{b}_1}^2 (z_{11} - z_{12})^2 + s_{\hat{a}_2}^2 s_{\hat{b}_2}^2 (z_{11}z_{21} - z_{12}z_{22})^2 \\ &+ 2s_{\hat{a}_1\hat{a}_2} s_{\hat{b}_2}^2 (z_{11}z_{21} - z_{12}z_{22})(z_{21} - z_{22}) + 2s_{\hat{b}_1\hat{b}_2} s_{\hat{a}_2}^2 (z_{11}z_{21} - z_{12}z_{22})(z_{11} - z_{12}) \\ &+ 2s_{\hat{a}_1\hat{a}_2} s_{\hat{b}_1\hat{b}_2} (z_{11} - z_{12})(z_{21} - z_{22}) \end{aligned}$$

Therefore, the first-order test statistic is $Z = \frac{f(\hat{\boldsymbol{\theta}})}{\sqrt{\mathbf{D}'\hat{\Sigma}(\hat{\boldsymbol{\theta}})\mathbf{D}}}$ and the second-order test statistic is $Z = \frac{f(\hat{\boldsymbol{\theta}})}{\sqrt{\mathbf{D}'\hat{\Sigma}(\hat{\boldsymbol{\theta}})\mathbf{D} + \frac{1}{2}tr\{(\mathbf{H}\hat{\Sigma}(\hat{\boldsymbol{\theta}}))^2\}}}$ accordingly.

Scenario 5: a and b are both moderated by Z

The null hypothesis is $H_0 : (a_1b_2 + a_2b_1)(z_1 - z_2) + a_2b_2(z_1^2 - z_2^2) = 0$ where z_1 and z_2 are two values of the moderator variable Z . For this scenario, we have

$$\hat{\boldsymbol{\theta}} = [\hat{a}_1 \ \hat{a}_2 \ \hat{b}_1 \ \hat{b}_2]$$

$$f(\hat{\boldsymbol{\theta}}) = (\hat{a}_1\hat{b}_2 + \hat{a}_2\hat{b}_1)(z_1 - z_2) + \hat{a}_2\hat{b}_2(z_1^2 - z_2^2)$$

$$\hat{\Sigma}(\hat{\boldsymbol{\theta}}) = \begin{bmatrix} s_{\hat{a}_1}^2 & s_{\hat{a}_1\hat{a}_2} & 0 & 0 \\ s_{\hat{a}_1\hat{a}_2} & s_{\hat{a}_2}^2 & 0 & 0 \\ 0 & 0 & s_{\hat{b}_1}^2 & s_{\hat{b}_1\hat{b}_2} \\ 0 & 0 & s_{\hat{b}_1\hat{b}_2} & s_{\hat{b}_2}^2 \end{bmatrix}$$

and thus

$$\begin{aligned} \mathbf{D}'\hat{\Sigma}(\hat{\boldsymbol{\theta}})\mathbf{D} &= (s_{\hat{b}_1}^2 \hat{a}_2^2 + s_{\hat{a}_1}^2 \hat{b}_2^2)(z_1 - z_2)^2 + 2(\hat{a}_1\hat{a}_2s_{\hat{b}_1\hat{b}_2} + \hat{b}_1\hat{b}_2s_{\hat{a}_1\hat{a}_2})(z_1 - z_2)^2 \\ &+ 2(\hat{a}_2^2s_{\hat{b}_1\hat{b}_2} + \hat{b}_2^2s_{\hat{a}_1\hat{a}_2})(z_1 - z_2)(z_1^2 - z_2^2) \\ &+ s_{\hat{b}_2}^2(\hat{a}_1(z_1 - z_2) + \hat{a}_2(z_1^2 - z_2^2))^2 + s_{\hat{a}_2}^2(\hat{b}_1(z_1 - z_2) + \hat{b}_2(z_1^2 - z_2^2))^2 \end{aligned}$$

$$\begin{aligned} \frac{1}{2}tr\{(\mathbf{H}\hat{\Sigma}(\hat{\boldsymbol{\theta}}))^2\} &= s_{\hat{a}_1}^2 s_{\hat{b}_2}^2 (z_1 - z_2)^2 + s_{\hat{a}_2}^2 s_{\hat{b}_1}^2 (z_1 - z_2)^2 + s_{\hat{a}_2}^2 s_{\hat{b}_2}^2 (z_1^2 - z_2^2)^2 \\ &+ 2s_{\hat{a}_1\hat{a}_2} s_{\hat{b}_1\hat{b}_2} (z_1 - z_2)^2 + 2(s_{\hat{a}_1\hat{a}_2} s_{\hat{b}_2}^2 + s_{\hat{b}_1\hat{b}_2} s_{\hat{a}_2}^2)(z_1^2 - z_2^2)(z_1 - z_2) \end{aligned}$$

Therefore, the first-order test statistic is $Z = \frac{f(\hat{\boldsymbol{\theta}})}{\sqrt{\mathbf{D}'\hat{\Sigma}(\hat{\boldsymbol{\theta}})\mathbf{D}}}$ and the second-order test statistic is

$Z = \frac{f(\hat{\theta})}{\sqrt{\mathbf{D}'\hat{\Sigma}(\hat{\theta})\mathbf{D} + \frac{1}{2}\text{tr}\{(\mathbf{H}\hat{\Sigma}(\hat{\theta}))^2\}}}$ accordingly.

Scenario 6: a is moderated by Z_1 and Z and b is moderated by Z_2 and Z

The null hypothesis H_0 is

$$\begin{aligned} a_1b_2(z_{21} - z_{22}) + a_2b_1(z_{11} - z_{12}) + a_2b_2(z_{11}z_{21} - z_{12}z_{22}) + (a_1b_3 + a_3b_1)(z_1 - z_2) \\ + a_2b_3(z_{11}z_1 - z_{12}z_2) + a_3b_2(z_1z_{21} - z_2z_{22}) + a_3b_3(z_1^2 - z_2^2) = 0 \end{aligned}$$

where z_{11} and z_{12} are two values of the moderator variable Z_1 , z_{21} and z_{22} are two corresponding values of the moderator variable Z_2 , and z_1 and z_2 are two values of the moderator variable Z .

The first-order and second-order variance estimates of the point estimate $f(\hat{\theta})$ can be derived by $\mathbf{D}'\hat{\Sigma}(\hat{\theta})\mathbf{D}$ and $\mathbf{D}'\hat{\Sigma}(\hat{\theta})\mathbf{D} + \frac{1}{2}\text{tr}\{(\mathbf{H}\hat{\Sigma}(\hat{\theta}))^2\}$. Here, we have

$$\hat{\theta} = [\hat{a}_1 \hat{a}_2 \hat{a}_3 \hat{b}_1 \hat{b}_2 \hat{b}_3]$$

$$\begin{aligned} f(\hat{\theta}) = \hat{a}_1\hat{b}_2(z_{21} - z_{22}) + \hat{a}_2\hat{b}_1(z_{11} - z_{12}) + \hat{a}_2\hat{b}_2(z_{11}z_{21} - z_{12}z_{22}) + (\hat{a}_1\hat{b}_3 + \hat{a}_3\hat{b}_1)(z_1 - z_2) \\ + \hat{a}_2\hat{b}_3(z_{11}z_1 - z_{12}z_2) + \hat{a}_3\hat{b}_2(z_1z_{21} - z_2z_{22}) + \hat{a}_3\hat{b}_3(z_1^2 - z_2^2) \end{aligned}$$

$$\hat{\Sigma}(\hat{\theta}) = \begin{bmatrix} s_{\hat{a}_1}^2 & s_{\hat{a}_1\hat{a}_2} & s_{\hat{a}_1\hat{a}_3} & 0 & 0 & 0 \\ s_{\hat{a}_1\hat{a}_2} & s_{\hat{a}_2}^2 & s_{\hat{a}_2\hat{a}_3} & 0 & 0 & 0 \\ s_{\hat{a}_1\hat{a}_3} & s_{\hat{a}_2\hat{a}_3} & s_{\hat{a}_3}^2 & 0 & 0 & 0 \\ 0 & 0 & 0 & s_{\hat{b}_1}^2 & s_{\hat{b}_1\hat{b}_2} & s_{\hat{b}_1\hat{b}_3} \\ 0 & 0 & 0 & s_{\hat{b}_1\hat{b}_2} & s_{\hat{b}_2}^2 & s_{\hat{b}_2\hat{b}_3} \\ 0 & 0 & 0 & s_{\hat{b}_1\hat{b}_3} & s_{\hat{b}_2\hat{b}_3} & s_{\hat{b}_3}^2 \end{bmatrix}$$

$$D(\hat{\theta}) = \begin{bmatrix} \hat{b}_2(z_{21} - z_{22}) + \hat{b}_3(z_1 - z_2) \\ \hat{b}_1(z_{11} - z_{12}) + \hat{b}_2(z_{11}z_{21} - z_{12}z_{22}) + \hat{b}_3(z_{11}z_1 - z_{12}z_2) \\ \hat{b}_1(z_1 - z_2) + \hat{b}_2(z_1z_{21} - z_2z_{22}) + \hat{b}_3(z_1^2 - z_2^2) \\ \hat{a}_2(z_{11} - z_{12}) + \hat{a}_3(z_1 - z_2) \\ \hat{a}_1(z_{21} - z_{22}) + \hat{a}_2(z_{11}z_{21} - z_{12}z_{22}) + \hat{a}_3(z_1z_{21} - z_2z_{22}) \\ \hat{a}_1(z_1 - z_2) + \hat{a}_2(z_{11}z_1 - z_{12}z_2) + \hat{a}_3(z_1^2 - z_2^2) \end{bmatrix}$$

and

$$H(\hat{\theta}) = \begin{bmatrix} 0 & 0 & 0 & 0 & z_{21} - z_{22} & z_1 - z_2 \\ 0 & 0 & 0 & z_{11} - z_{12} & z_{11}z_{21} - z_{12}z_{22} & z_1z_{11} - z_2z_{12} \\ 0 & 0 & 0 & z_1 - z_2 & z_1z_{21} - z_2z_{22} & z_1^2 - z_2^2 \\ 0 & z_{11} - z_{12} & z_1 - z_2 & 0 & 0 & 0 \\ z_{21} - z_{22} & z_{11}z_{21} - z_{12}z_{22} & z_1z_{21} - z_2z_{22} & 0 & 0 & 0 \\ z_1 - z_2 & z_1z_{11} - z_2z_{12} & z_1^2 - z_2^2 & 0 & 0 & 0 \end{bmatrix}$$

Since the final forms of $D'\hat{\Sigma}(\hat{\theta})D$ and $D'\hat{\Sigma}(\hat{\theta})D + \frac{1}{2}tr\{(H\hat{\Sigma}(\hat{\theta}))^2\}$ are very cumbersome, we do not display them here for the sake of saving space. The first-order test statistic is $Z = \frac{f(\hat{\theta})}{\sqrt{D'\hat{\Sigma}(\hat{\theta})D}}$ and the second-order test statistic is $Z = \frac{f(\hat{\theta})}{\sqrt{D'\hat{\Sigma}(\hat{\theta})D + \frac{1}{2}tr\{(H\hat{\Sigma}(\hat{\theta}))^2\}}}$ accordingly.

References

Preacher, K. J., Rucker, D. D., & Hayes, A. F. (2007). Addressing moderated mediation hypotheses: Theory, methods, and prescriptions. *Multivariate Behavioral Research*, 42 (1), 185 – 227.