

Models for Group Comparisons – Summary

Richard Williams, University of Notre Dame, <https://www3.nd.edu/~rwilliam/>

Last revised February 20, 2015

Since we are estimating and comparing several models, it will be helpful to list several of them all in one place. This handout summarizes how to do group comparisons by running separate models for each group and by using interaction terms. We won't go through this handout separately, but it may help you to keep everything straight.

I. **Pooled (Constrained) Model.** Blacks and whites are combined into a single analysis, hence the coefficient estimates are constrained to be the same for both racial groups, i.e. the intercepts and the effect of education and job experience are the same for both groups.

```
. use "https://www3.nd.edu/~rwilliam/statafiles/blwh.dta"
. reg income educ jobexp
```

| Source | SS | df | MS | Number of obs = 500 | | |
|----------|------------|-----|------------|---------------------|---|---------|
| Model | 32798.4018 | 2 | 16399.2009 | F(2, 497) | = | 1103.96 |
| Residual | 7382.84742 | 497 | 14.8548238 | Prob > F | = | 0.0000 |
| ----- | | | | R-squared | = | 0.8163 |
| Total | 40181.2493 | 499 | 80.5235456 | Adj R-squared | = | 0.8155 |
| ----- | | | | Root MSE | = | 3.8542 |

| income | Coef. | Std. Err. | t | P> t | [95% Conf. Interval] | |
|--------|-----------|-----------|-------|-------|----------------------|-----------|
| educ | 1.94512 | .0436998 | 44.51 | 0.000 | 1.859261 | 2.03098 |
| jobexp | .7082212 | .0343672 | 20.61 | 0.000 | .6406983 | .775744 |
| _cons | -7.382935 | .8027781 | -9.20 | 0.000 | -8.960192 | -5.805678 |

II. **Unconstrained Models: Separate Models for each group.** Here, we estimate separate models, first for whites, then blacks. This makes it possible for the intercepts and slope coefficients to freely differ across populations. This is equivalent to Model IIIC.

Whites:

```
. reg income educ jobexp if black == 0
```

| Source | SS | df | MS | Number of obs = 400 | | |
|----------|------------|-----|------------|---------------------|---|--------|
| Model | 18361.9894 | 2 | 9180.99472 | F(2, 397) | = | 620.07 |
| Residual | 5878.16991 | 397 | 14.8064733 | Prob > F | = | 0.0000 |
| ----- | | | | R-squared | = | 0.7575 |
| Total | 24240.1594 | 399 | 60.7522791 | Adj R-squared | = | 0.7563 |
| ----- | | | | Root MSE | = | 3.8479 |

| income | Coef. | Std. Err. | t | P> t | [95% Conf. Interval] | |
|--------|-----------|-----------|-------|-------|----------------------|----------|
| educ | 1.893338 | .0562591 | 33.65 | 0.000 | 1.782735 | 2.003941 |
| jobexp | .722255 | .0412236 | 17.52 | 0.000 | .6412111 | .8032988 |
| _cons | -6.461189 | 1.089219 | -5.93 | 0.000 | -8.602547 | -4.31983 |

Blacks:

```
. reg income educ jobexp if black == 1
```

| Source | SS | df | MS | Number of obs = 100 | | |
|----------|------------|----|------------|---------------------|---|--------|
| Model | 4924.27286 | 2 | 2462.13643 | F(2, 97) | = | 267.80 |
| Residual | 891.81705 | 97 | 9.19399021 | Prob > F | = | 0.0000 |
| ----- | | | | R-squared | = | 0.8467 |
| ----- | | | | Adj R-squared | = | 0.8435 |
| Total | 5816.08991 | 99 | 58.748383 | Root MSE | = | 3.0322 |

| income | Coef. | Std. Err. | t | P> t | [95% Conf. Interval] | |
|--------|----------|-----------|-------|-------|----------------------|-----------|
| educ | 1.677949 | .0725479 | 23.13 | 0.000 | 1.533962 | 1.821936 |
| jobexp | .421975 | .0581021 | 7.26 | 0.000 | .3066585 | .5372915 |
| _cons | -3.0512 | 1.154604 | -2.64 | 0.010 | -5.342771 | -.7596302 |

Doing an incremental F Test (see earlier notes for details):

$$F_{K+1, N_1+N_2-2K-2} = \frac{(SSE_c - SSE_u) * (N_1 + N_2 - 2K - 2)}{SSE_u * (K + 1)} = \frac{(7383 - 6770) * 494}{6770 * 3} = 14.91$$

III. Unconstrained and Partially Unconstrained Models: Dummy Variables and Interaction Effects. In this approach, interaction effects and dummy variables are used to allow for racial differences in parameter effects. Interaction effects allow more flexibility in model specification.

IIIa. Intercepts vary across groups, all other parameters the same. To do this, you regress Y on the IVs and include a dummy variable for race.

```
. reg income educ jobexp i.black
```

| Source | SS | df | MS | Number of obs = 500 | | |
|----------|------------|-----|------------|---------------------|---|--------|
| Model | 33206.4588 | 3 | 11068.8196 | F(3, 496) | = | 787.14 |
| Residual | 6974.79047 | 496 | 14.0620776 | Prob > F | = | 0.0000 |
| ----- | | | | R-squared | = | 0.8264 |
| ----- | | | | Adj R-squared | = | 0.8254 |
| Total | 40181.2493 | 499 | 80.5235456 | Root MSE | = | 3.7499 |

| income | Coef. | Std. Err. | t | P> t | [95% Conf. Interval] | |
|---------|-----------------|-----------|--------------|-------|----------------------|-----------|
| educ | 1.840407 | .0467507 | 39.37 | 0.000 | 1.748553 | 1.932261 |
| jobexp | .6514259 | .0350604 | 18.58 | 0.000 | .5825406 | .7203111 |
| 1.black | -2.55136 | .4736266 | -5.39 | 0.000 | -3.481921 | -1.620798 |
| _cons | -4.72676 | .9236842 | -5.12 | 0.000 | -6.541576 | -2.911943 |

```
. * Wald test of differences in intercepts
. testparm i.black
```

```
( 1) 1.black = 0
```

```
F( 1, 496) = 29.02
Prob > F = 0.0000
```

IIIb. Intercepts and some slopes vary across groups, other slopes are the same.
 Regress Y on the IVs, a dummy variable for race, and (in this example) one interaction term:

```
. reg income educ jobexp i.black i.black#c.jobexp
```

| Source | SS | df | MS | Number of obs = 500 | | |
|----------|------------|-----|------------|---------------------|---|--------|
| Model | 33352.2559 | 4 | 8338.06397 | F(4, 495) | = | 604.39 |
| Residual | 6828.99339 | 495 | 13.7959462 | Prob > F | = | 0.0000 |
| | | | | R-squared | = | 0.8300 |
| | | | | Adj R-squared | = | 0.8287 |
| Total | 40181.2493 | 499 | 80.5235456 | Root MSE | = | 3.7143 |

| income | Coef. | Std. Err. | t | P> t | [95% Conf. Interval] | |
|----------------|-----------|-----------|-------|-------|----------------------|-----------|
| educ | 1.834776 | .0463385 | 39.60 | 0.000 | 1.743732 | 1.925821 |
| jobexp | .7128145 | .0395293 | 18.03 | 0.000 | .6351486 | .7904805 |
| 1.black | .4686862 | 1.040728 | 0.45 | 0.653 | -1.576103 | 2.513475 |
| black#c.jobexp | | | | | | |
| 1 | -.2556117 | .0786289 | -3.25 | 0.001 | -.4100993 | -.1011242 |
| _cons | -5.514076 | .9464143 | -5.83 | 0.000 | -7.373561 | -3.654592 |

```
. * Wald test of racial difference in effect of jobexp
. testparm i.black#c.jobexp
```

```
( 1) 1.black#c.jobexp = 0
```

```
      F( 1, 495) = 10.57
      Prob > F = 0.0012
```

IIIc. Intercepts and all slopes free to vary across groups (totally unconstrained). Equivalent to Model II, where we estimated separate models for each group.

```
. reg income educ jobexp i.black i.black#c.educ i.black#c.jobexp
```

| Source | SS | df | MS | Number of obs = 500 | | |
|----------|------------|-----|------------|---------------------|---|--------|
| Model | 33411.2623 | 5 | 6682.25246 | F(5, 494) | = | 487.60 |
| Residual | 6769.98696 | 494 | 13.7044271 | Prob > F | = | 0.0000 |
| | | | | R-squared | = | 0.8315 |
| | | | | Adj R-squared | = | 0.8298 |
| Total | 40181.2493 | 499 | 80.5235456 | Root MSE | = | 3.7019 |

| income | Coef. | Std. Err. | t | P> t | [95% Conf. Interval] | |
|----------------|-----------|-----------|-------|-------|----------------------|-----------|
| educ | 1.893338 | .054125 | 34.98 | 0.000 | 1.786994 | 1.999681 |
| jobexp | .722255 | .0396598 | 18.21 | 0.000 | .6443322 | .8001777 |
| 1.black | 3.409988 | 1.756477 | 1.94 | 0.053 | -.0410984 | 6.861075 |
| black#c.educ | | | | | | |
| 1 | -.2153886 | .1038015 | -2.08 | 0.039 | -.4193354 | -.0114418 |
| black#c.jobexp | | | | | | |
| 1 | -.3002799 | .0812705 | -3.69 | 0.000 | -.4599584 | -.1406015 |
| _cons | -6.461189 | 1.0479 | -6.17 | 0.000 | -8.520079 | -4.402298 |

```
. * Wald test of differences in slopes
```

```
. testparm i.black#c.educ i.black#c.jobexp
```

```
( 1) 1.black#c.educ = 0
( 2) 1.black#c.jobexp = 0
```

```
F( 2, 494) = 7.47
Prob > F = 0.0006
```

```
. * Wald test of any differences across groups, including intercepts
```

```
. testparm i.black i.black#c.educ i.black#c.jobexp
```

```
( 1) 1.black = 0
( 2) 1.black#c.educ = 0
( 3) 1.black#c.jobexp = 0
```

```
F( 3, 494) = 14.91
Prob > F = 0.0000
```

Note that $N_u = 500$, $SSE_u = 6770$, $DFE_u = 494$. These are the exact same numbers we got for the incremental F test using the Model II procedure of estimating separate models for each racial group. Further, the regression coefficients estimated under this approach can easily be converted to the betas estimated under the previous approach, and vice versa.