

Sociology 63993, Exam 2 Answer Key [DRAFT] March 27, 2015

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I. True-False. (20 points) Indicate whether the following statements are true or false. If false, briefly explain why.

1. A researcher has inadvertently included an extraneous variable in her model. Unfortunately, increasing the sample size will not help to reduce the problems this creates.

False. Extraneous variables increase standard errors, while a larger sample size reduces standard errors. The larger your sample is, the more you can afford to include extraneous variables – which you may want to do just to explicitly show that their effects are 0.

2. A researcher regresses income on the respondent's gender, years of education, IQ, and mother's education (i.e. the number of years of education the respondent's mother had). The estimated effect of mother's education is 0 and is statistically insignificant. This means that, in terms of their own income, respondents gain no benefit from having a better educated mother.

False (or at least not necessarily true). Mother's education could still have indirect effects, e.g. more education for the mother could result in more education for the child which in turn results in a higher income.

3. Personal Fulfillment (measured on a 200 point scale) is regressed on Income, Female, and Female*Income. All terms are positive and statistically significant. The coefficient for Female is +12. This means that, whenever a man and a woman have equal incomes, the woman is expected to score 12 points higher than the man on Personal Fulfillment.

False. Because there is an interaction term, the expected difference between an otherwise identical man and women is only 12 when income = 0. At other values of income, the expected difference between the man and the woman will either be greater than 12 or less than 12.

4. A researcher believes that X2 and X3 are positively correlated only because X1 is a common cause of both, i.e. X2 does not directly or indirectly affect X3, nor does X3 directly or indirectly affect X2. Therefore knowledge of X2 will be of no use to her for predicting X3.

False. Because X2 and X3 are positively correlated, you know that those who have higher values on X2 will also tend to have higher values on X3. Just because a relationship is spurious or non-causal does not mean that it can't be useful for making predictions.

5. A researcher hypothesizes that income positively affects the self-image of men but has a negative effect on the self-image of women. She gets

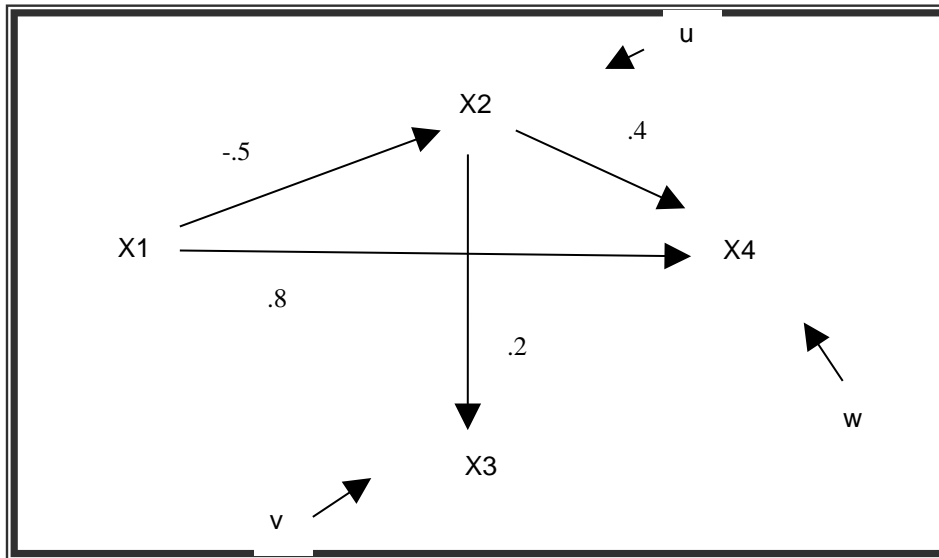
$$\hat{\beta}_{Income} = 4$$

$$\hat{\beta}_{Female} = 0$$

$$\hat{\beta}_{Income * Female} = -4$$

Female = 1 if female, 0 if male. The T values for Income and for the interaction term are both highly significant. The evidence supports the researcher's hypothesis.

False. These results say that the effect of income is 0 for women, not negative.



II. Path Analysis/Model specification (25 pts). A sociologist believes that the following model describes the relationship between X1, X2, X3, and X4. All her variables are in standardized form. The estimated value of each path in her model is included in the diagram.

a. (5 pts) Write out the structural equation for each endogenous variable, using both the names for the paths (e.g. β_{42}) and the estimated value of the path coefficient.

$$X_2 = \beta_{21}X_1 + u = -.5X_1 + u$$

$$X_3 = \beta_{32}X_2 + v = .2X_2 + v$$

$$X_4 = \beta_{41}X_1 + \beta_{42}X_2 + w = .8X_1 + .4X_2 + w$$

b. (10 pts) Part of the correlation matrix is shown below. Determine the complete correlation matrix. (Remember, variables are standardized. You can use either normal equations or Sewell Wright, but you might want to use both as a double-check.)

	x1	x2	x3	x4
x1	1.0000			
x2	-0.5000	1.0000		
x3	?	?	1.0000	
x4	.?	?	?	1.0000

Here is the uncensored printout:

```

. corr
(obs=100)

```

	x1	x2	x3	x4
x1	1.0000			
x2	-0.5000	1.0000		
x3	-0.1000	0.2000	1.0000	
x4	0.6000	0.0000	0.0000	1.0000

To compute by hand,

$$\rho_{31} = \beta_{31} + \beta_{21}\beta_{32} = 0 + (-.5*.2) = -.10$$

$$\rho_{32} = \beta_{32} + \beta_{31}\beta_{21} = .2 + (0*-.5) = .2$$

$$\rho_{41} = \beta_{41} + \beta_{21}\beta_{42} + \beta_{31}\beta_{43} + \beta_{21}\beta_{32}\beta_{43} = .8 + (-.5*.4) + (0*0) + (-.5*2*0) = .60$$

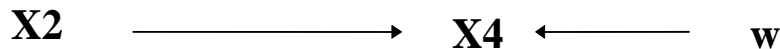
$$\rho_{42} = \beta_{42} + \beta_{32}\beta_{43} + \beta_{41}\beta_{21} + \beta_{43}\beta_{31}\beta_{21} = .4 + (.2*0) + (.8*-.5) + (0*0*-.5) = 0$$

$$\rho_{43} = \beta_{43} + \beta_{41}\beta_{31} + \beta_{42}\beta_{32} + \beta_{41}\beta_{21}\beta_{32} + \beta_{42}\beta_{21}\beta_{31} = 0 + (.8*0) + (.4*.2) + (.8*-.5*.2) + (.4*-.5*0) = 0$$

c. (5 pts) Decompose the correlation between X1 and X4 into

- Correlation due to direct effects
.8
- Correlation due to indirect effects
-.2
- Correlation due to common causes
0

d. (5 pts) Suppose the above model is correct, but instead the researcher believed in and estimated the following model:



What conclusions would the researcher likely draw? In particular, what would the researcher conclude about the effect of changes in X2 on X4? Discuss the consequences of this mis-specification, and in what ways, if any, the results would be misleading. Why would she make these mistakes?

Since this is a bivariate regression the slope coefficient is the same as the correlation, which in this case is 0. The researcher will therefore conclude that changes in X2 have no effect on X4, when in reality the effect of X2 is .4, meaning that increases in X2 tend to produce increases in X4. This could be a very serious mistake if it led the researcher to make policy decisions based on the belief that X2 had no effect on X4. There is omitted variable bias because X1 should be in the model but it is not. The fact that suppressor effects are also present further leads to the incorrect conclusion.

To confirm above results using Stata commands,

```
. * Problem II
. clear all
. matrix input corr = (1,-.5,-.1,.6\-.5,1,.2,0\-.1,.2,1,0\-.6,0,0,1)
. corr2data x1 x2 x3 x4, n(100) corr(corr)
(obs 100)

. * Confirm results
```

. corr
(obs=100)

	x1	x2	x3	x4
x1	1.0000			
x2	-0.5000	1.0000		
x3	-0.1000	0.2000	1.0000	
x4	0.6000	0.0000	0.0000	1.0000

. reg x2 x1

Source	SS	df	MS	Number of obs =	100
Model	24.7500004	1	24.7500004	F(1, 98) =	32.67
Residual	74.2500018	98	.757653079	Prob > F =	0.0000
				R-squared =	0.2500
				Adj R-squared =	0.2423
Total	99.0000022	99	1.00000002	Root MSE =	.87043

x2	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
x1	-.5	.0874818	-5.72	0.000	-.6736047 -.3263953
_cons	-1.73e-09	.0870433	-0.00	1.000	-.1727345 .1727345

. reg x3 x2 x1

Source	SS	df	MS	Number of obs =	100
Model	3.95999991	2	1.97999995	F(2, 97) =	2.02
Residual	95.0400006	97	.97979382	Prob > F =	0.1381
				R-squared =	0.0400
				Adj R-squared =	0.0202
Total	99.0000005	99	1	Root MSE =	.98985

x3	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
x2	.2	.1148733	1.74	0.085	-.0279917 .4279917
x1	9.07e-09	.1148733	0.00	1.000	-.2279917 .2279917
_cons	-6.59e-09	.0989845	-0.00	1.000	-.1964569 .1964569

. reg x4 x3 x2 x1

Source	SS	df	MS	Number of obs =	100
Model	47.5199999	3	15.84	F(3, 96) =	29.54
Residual	51.4799996	96	.536249996	Prob > F =	0.0000
				R-squared =	0.4800
				Adj R-squared =	0.4638
Total	98.9999996	99	.999999996	Root MSE =	.73229

x4	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
x3	8.81e-09	.0751157	0.00	1.000	-.1491034 .1491034
x2	.4	.0863013	4.63	0.000	.2286932 .5713067
x1	.8	.0849837	9.41	0.000	.6313088 .9686912
_cons	-5.44e-09	.0732291	-0.00	1.000	-.1453586 .1453586

. pathreg (x2 x1) (x3 x2 x1) (x4 x1 x2 x3)

```

-----
      x2 |           Coef.   Std. Err.       t   P>|t|                Beta
-----+-----
      x1 |             -.5   .0874818      -5.72  0.000               -.5
     _cons |    -1.73e-09   .0870433      -0.00  1.000                .
-----
          n = 100   R2 = 0.2500  sqrt(1 - R2) = 0.8660

```

```

-----
      x3 |           Coef.   Std. Err.       t   P>|t|                Beta
-----+-----
      x2 |             .2   .1148733       1.74  0.085                .2
      x1 |    9.07e-09   .1148733       0.00  1.000              9.07e-09
     _cons |   -6.59e-09   .0989845      -0.00  1.000                .
-----
          n = 100   R2 = 0.0400  sqrt(1 - R2) = 0.9798

```

```

-----
      x4 |           Coef.   Std. Err.       t   P>|t|                Beta
-----+-----
      x1 |             .8   .0849837       9.41  0.000                .8
      x2 |             .4   .0863013       4.63  0.000                .4
      x3 |    8.81e-09   .0751157       0.00  1.000              8.81e-09
     _cons |   -5.44e-09   .0732291      -0.00  1.000                .
-----
          n = 100   R2 = 0.4800  sqrt(1 - R2) = 0.7211

```

```

. sem (x2 <- x1) (x3 <- x2 x1) (x4 <- x1 x2 x3)

```

```

Endogenous variables

```

```

Observed:  x2 x3 x4

```

```

Exogenous variables

```

```

Observed:  x1

```

```

Fitting target model:

```

```

Iteration 0:  log likelihood = -516.44382
Iteration 1:  log likelihood = -516.44382

```

```

Structural equation model           Number of obs       =           100
Estimation method = ml
Log likelihood       = -516.44382

```

```

-----
                        |           Coef.    OIM              [95% Conf. Interval]
                        |           Std. Err.    z     P>|z|
-----+-----
Structural
  x2 <-
      x1 |             -.5   .0866025      -5.77  0.000       -.6697379   -.3302621
     _cons |    -1.73e-09   .0861684      -0.00  1.000       -.1688887   .1688887
-----
  x3 <-
      x2 |             .2   .1131371       1.77  0.077       -.0217446   .4217446
      x1 |    9.07e-09   .1131371       0.00  1.000       -.2217446   .2217446
     _cons |   -6.59e-09   .0974885      -0.00  1.000       -.1910739   .1910739
-----
  x4 <-

```

x2	.4	.0845577	4.73	0.000	.23427	.56573
x3	8.81e-09	.073598	0.00	1.000	-.1442494	.1442495
x1	.8	.0832666	9.61	0.000	.6368004	.9631996
_cons	-5.44e-09	.0717496	-0.00	1.000	-.1406266	.1406266

var(e.x2)	.7425	.1050054			.5627537	.9796581
var(e.x3)	.9504	.1344069			.7203248	1.253962
var(e.x4)	.5148	.0728037			.3901759	.6792296

LR test of model vs. saturated: chi2(0) = 0.00, Prob > chi2 = .

. estat teffects

Direct effects

	Coef.	OIM Std. Err.	z	P> z	[95% Conf. Interval]	

Structural						
x2 <-						
x1	-.5	.0866025	-5.77	0.000	-.6697379	-.3302621

x3 <-						
x2	.2	.1131371	1.77	0.077	-.0217446	.4217446
x1	9.07e-09	.1131371	0.00	1.000	-.2217446	.2217446

x4 <-						
x2	.4	.0845577	4.73	0.000	.23427	.56573
x3	8.81e-09	.073598	0.00	1.000	-.1442494	.1442495
x1	.8	.0832666	9.61	0.000	.6368004	.9631996

Indirect effects

	Coef.	OIM Std. Err.	z	P> z	[95% Conf. Interval]	

Structural						
x2 <-						
x1	0	(no path)				

x3 <-						
x2	0	(no path)				
x1	-.1	.0591608	-1.69	0.091	-.215953	.015953

x4 <-						
x2	1.76e-09	(constrained)				
x3	0	(no path)				
x1	-.2	.0541603	-3.69	0.000	-.3061522	-.0938479

Total effects

	Coef.	OIM Std. Err.	z	P> z	[95% Conf. Interval]	

Structural						
x2 <-						
x1	-.5	.0866025	-5.77	0.000	-.6697379	-.3302621

```

x3 <-
  x2 |          .2   .1131371    1.77   0.077   -.0217446   .4217446
  x1 |         -.1   .0994987   -1.01   0.315   -.2950139   .095014
-----
x4 <-
  x2 |          .4   .0845577    4.73   0.000    .23427    .56573
  x3 |  8.81e-09   .073598    0.00   1.000   -.1442494   .1442495
  x1 |          .6    .08    7.50   0.000    .4432029   .7567971
-----

. * Erroneous model
. reg x4 x2

      Source |           SS          df           MS          Number of obs =      100
-----+-----+-----+-----+-----+-----
      Model |             0            1             0          F( 1, 98) =      0.00
      Residual | 98.9999996          98    1.01020408        Prob > F      = 1.0000
-----+-----+-----+-----+-----
      Total | 98.9999996          99    .999999996        R-squared      = 0.0000
                                          Adj R-squared  = -0.0102
                                          Root MSE      = 1.0051

-----+-----+-----+-----+-----+-----
      x4 |           Coef.    Std. Err.      t    P>|t|    [95% Conf. Interval]
-----+-----+-----+-----+-----+-----
      x2 |  5.20e-09    .1010153      0.00   1.000    - .2004615    .2004615
      _cons | -6.20e-09    .1005089     -0.00   1.000    - .1994567    .1994567
-----+-----+-----+-----+-----+-----

```

III. Group comparisons (25 points). Yik Yak is an anonymous social media app. It allows users to share posts with others who are within a few miles of them, and has become very popular on college campuses, including Notre Dame. However, the presence of racist and sexist posts on Yik Yak has generated great controversy, leading, for example, to 150 Notre Dame faculty signing a letter denouncing some of the postings (http://www.southbendtribune.com/news/notre-dame-profs-students-respond-to-racist-yik-yak-posts/article_b1f2d81a-a311-11e4-ae3a-1b3bd646ee0d.html?_dc=479429358849.30194). Potential investors in Yik Yak are worried that such controversies could undermine the financial prospects for the application. They want to assess how serious the concerns about offensive posts are. They have therefore conducted a study of 2000 randomly selected college students. Participants were asked to use Yik Yak for a month, and then provide answers to the following questions:

Variable	Description
yikyak	How likely is the respondent to use yikyak in the future? Scores range from a low of 1 (definitely will not use) to a high of 150 (certain to use).
offensive	How offensive did the user find yikyak? The original scale ranged from a low of 1(not offensive at all) to a high of 100 (extremely offensive). The scale has been centered so that a score of zero corresponds to an average score on the original measure.
female	Coded 1 if female, 0 if male
femXoffensive	female * offensive.

The results of the analysis are as follows:

. ttest yikyak, by(female)

Two-sample t test with equal variances

```
-----+-----
      Group |      Obs      Mean      Std. Err.      Std. Dev.      [95% Conf. Interval]
-----+-----
          0 |        924    63.28586    1.039956    31.61193    61.24491    65.32681
          1 |       1076    44.10861    .9983174    32.74725    42.14974    46.06748
-----+-----
combined |       2000    52.9685     .7515199    33.60899    51.49465    54.44234
-----+-----
      diff |              19.17725    1.445449              16.34251    22.012
-----+-----
      diff = mean(0) - mean(1)                                t = 13.2673
Ho: diff = 0                                                degrees of freedom = 1998

      Ha: diff < 0              Ha: diff != 0              Ha: diff > 0
Pr(T < t) = 1.0000          Pr(|T| > |t|) = 0.0000          Pr(T > t) = 0.0000
```

. nestreg: reg yikyak offensive female femXoffensive

Block 1: offensive

```
-----+-----
      Source |      SS      df      MS              Number of obs = 2000
-----+-----
      Model | 343159.58      1 343159.58          F( 1, 1998) = 358.06
      Residual | 1914839.55 1998  958.378155          Prob > F      = 0.0000
-----+-----
      Total | 2257999.13 1999 1129.56435          R-squared     = 0.1520
                                      Adj R-squared = 0.1516
                                      Root MSE     = 30.958
-----+-----
      yikyak |      Coef.      Std. Err.      t      P>|t|      [95% Conf. Interval]
-----+-----
      offensive | - .8789011    .0464473   -18.92   0.000   - .9699913   - .7878109
      _cons | 52.9685     .6922348    76.52   0.000    51.61092    54.32608
-----+-----
```

Block 2: female

```
-----+-----
      Source |      SS      df      MS              Number of obs = 2000
-----+-----
      Model | 452528.792      2 226264.396          F( 2, 1997) = 250.27
      Residual | 1805470.34 1997  904.091308          Prob > F      = 0.0000
-----+-----
      Total | 2257999.13 1999 1129.56435          R-squared     = 0.2004
                                      Adj R-squared = 0.1996
                                      Root MSE     = 30.068
-----+-----
      yikyak |      Coef.      Std. Err.      t      P>|t|      [95% Conf. Interval]
-----+-----
      offensive | - .7912462    .0458112   -17.27   0.000   - .8810889   - .7014035
      female | -15.06238    1.369469   -11.00   0.000   -17.74812   -12.37664
      _cons | 61.07206     .9974378    61.23   0.000    59.11593    63.02819
-----+-----
```


Block 3: femXoffensive

Source	SS	df	MS	Number of obs = 2000		
Model	465693.904	3	155231.301	F(3, 1996)	=	172.87
Residual	1792305.23	1996	897.948512	Prob > F	=	0.0000
				R-squared	=	0.2062
				Adj R-squared	=	0.2050
				Root MSE	=	29.966

yikyak	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
offensive	-.5990285	.0678563	-8.83	0.000	-.7321051	-.4659519
female	-15.21819	1.365415	-11.15	0.000	-17.89598	-12.54041
femXoffensive	-.3512047	.091722	-3.83	0.000	-.5310857	-.1713237
_cons	61.60986	1.003917	61.37	0.000	59.64103	63.5787

Block	F	Block df	Residual df	Pr > F	R2	Change in R2
1	358.06	1	1998	0.0000	0.1520	
2	120.97	1	1997	0.0000	0.2004	0.0484
3	14.66	1	1996	0.0001	0.2062	0.0058

. ttest offensive, by(female)

Two-sample t test with equal variances

Group	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
0	924	-2.797863	.478188	14.53566	-3.736325	-1.859401
1	1076	2.402631	.4514825	14.80973	1.516744	3.288518
combined	2000	2.53e-06	.3333394	14.90739	-.6537265	.6537316
diff		-5.200494	.6585821		-6.492073	-3.908914
diff = mean(0) - mean(1)				t =	-7.8965	
Ho: diff = 0				degrees of freedom =	1998	
Ha: diff < 0		Ha: diff != 0		Ha: diff > 0		
Pr(T < t) = 0.0000		Pr(T > t) = 0.0000		Pr(T > t) = 1.0000		

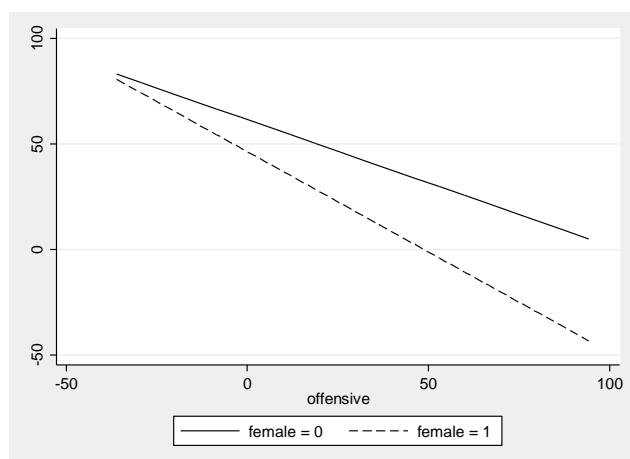
The initial t-test shows that women are significantly less likely to keep using Yik Yak in the future. Based on the remaining results, explain to Yik Yak’s backers why that is the case. When thinking about your answers, keep in mind the various reasons that two groups can differ on some outcome measure. Specifically, answer the following:

- a) (10 pts) The researchers estimate a series of models. Which of the models do you think is best, and why? What do these models tell us about how gender and the perceived offensiveness of Yik Yak affect the likelihood of using Yik Yak in the future? What ways (if any) do the determinants of support for Yik Yak differ by gender?

All of the coefficients in model 3 are statistically significant so (at least from a purely empirical standpoint) it is the best model. According to model 3, the more offensive someone finds Yik Yak, the less likely they are to continue using it. Further, this negative effect is significantly greater for women than it is men. For men, the estimated effect of offensive is -.599 but for women it is -.950.

This graph will also help to show the relationships. I use the mcp command, which I discuss more in my Categorical Data Analysis course. It shows that the differences between men and women are very small among those who don't find Yik Yak very offensive but then get greater and greater. (OK, I lied about the range of yikyak but it doesn't really matter – If I was more ambitious I could just add a constant to the scores and everything would be the same except the intercepts.)

```
. quietly reg yikyak offensive i.female i.female#c.offensive
. mcp offensive female, var1(20)
```



- b) (5 pts) According to your preferred model, how does the yikyak score of the “average” male compare to that of the “average” female?

Because offensive is centered, the “average” person has a score of 0 on it. When offensive equals 0, women score 15.22 points lower on yikyak than men do.

- c) (10 pts) The researchers then do one last t-test. What does this test tell us about how feelings on offensiveness differ by gender? What additional insights, if any, does this test give us as to why women are less supportive of Yik Yak?

Women outscore men by 5.2 points on the offensive scale. So, even if the effect of offensive did not vary by gender, women would have lower scores on yikyak. The fact that women find yikyak more offensive, and the fact that the effect of offensive is greater for women than it is men, both help to create the disparities between men and women in their likelihood of using Yik Yak in the future.

IV. Short answer. Answer *both* of the following questions. (15 points each, 30 points total.) In each of the following problems, a researcher runs through a sequence of commands. Explain why she didn't stop after the first command, i.e. explain what the purpose of each subsequent command was, what it told her, and why she did not run additional commands after the last one. If she had stopped after the first command, what would the consequences have been, i.e. in what ways would her conclusions have been incorrect or misleading? Include diagrams or scatterplots that describe the relationships if they have not already been provided in the problem.

1.

```
. reg liberalism ses
```

Source	SS	df	MS	Number of obs =	500
Model	567994.874	1	567994.874	F(1, 498) =	2764.60
Residual	102315.49	498	205.452792	Prob > F =	0.0000
				R-squared =	0.8474
				Adj R-squared =	0.8471
				Root MSE =	14.334
Total	670310.364	499	1343.30734		

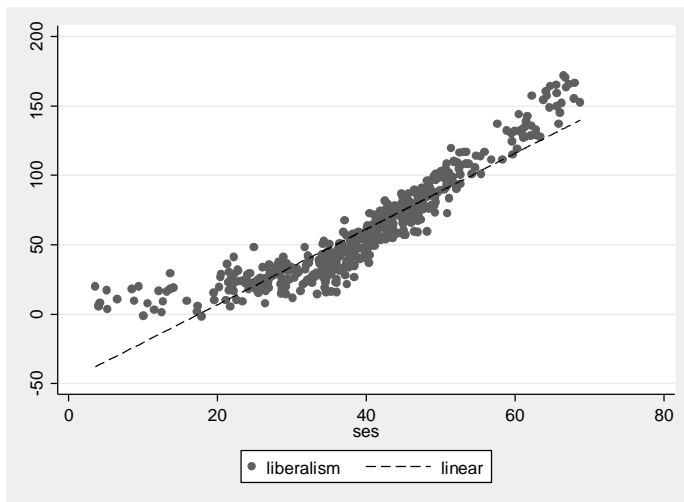
liberalism	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
ses	2.729271	.0519075	52.58	0.000	2.627286 2.831256
_cons	-47.85882	2.156255	-22.20	0.000	-52.0953 -43.62234

```
. predict linear
```

(option xb assumed; fitted values)

```
. label variable linear "linear"
```

```
. scatter liberalism ses || line linear ses, scheme(sj) sort
```



```
. mkspline seslow 36 seshi = ses
```

```
. reg liberalism seslow seshi
```

Source	SS	df	MS	Number of obs =	500
Model	632827.848	2	316413.924	F(2, 497) =	4195.50
Residual	37482.5161	497	75.4175374	Prob > F =	0.0000
				R-squared =	0.9441
				Adj R-squared =	0.9439
				Root MSE =	8.6843
Total	670310.364	499	1343.30734		

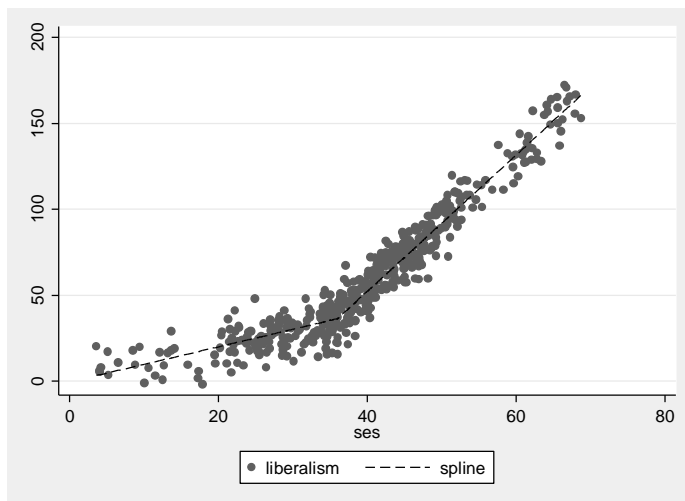
liberalism	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
seslow	1.022624	.0661605	15.46	0.000	.8926348 1.152612
seshi	3.965107	.0525898	75.40	0.000	3.861781 4.068433
_cons	-.424041	2.079451	-0.20	0.838	-4.509639 3.661557

```
. predict spline
```

```
(option xb assumed; fitted values)
```

```
. label variable spline "spline"
```

```
. scatter liberalism ses || line spline ses, scheme(sj) sort
```



The original model had a very high R^2 . However, the scatterplot suggested that SES initially only had a very small effect on liberalism, and after around $SES = 36$ the effect became much larger. Perhaps the researcher also had good reasons for thinking that 36 and above would have a much larger effect. She therefore used a spline function, which allowed the effect of SES to differ before and after $SES = 36$. Not only did this improve the R^2 , the graph of the predicted and observed values looked much better. As the first graph shows, a simple linear model would have overestimated or underestimated the effect of SES on liberalism at different values of SES.

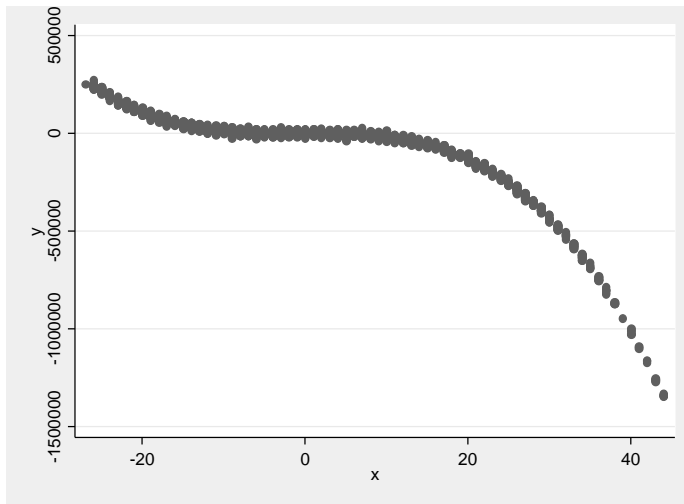
2.

. reg y x

Source	SS	df	MS	Number of obs =	2293
Model	5.9161e+13	1	5.9161e+13	F(1, 2291) =	4666.20
Residual	2.9047e+13	2291	1.2679e+10	Prob > F =	0.0000
Total	8.8208e+13	2292	3.8485e+10	R-squared =	0.6707
				Adj R-squared =	0.6706
				Root MSE =	1.1e+05

y	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
x	-9575.092	140.1721	-68.31	0.000	-9849.97 -9300.215
_cons	-37867.18	2351.439	-16.10	0.000	-42478.35 -33256.01

. scatter y x, scheme(sj)



. reg y x c.x#c.x c.x#c.x#c.x

Source	SS	df	MS	Number of obs =	2293
Model	8.7979e+13	3	2.9326e+13	F(3, 2289) =	.
Residual	2.2843e+11	2289	99793436.9	Prob > F =	0.0000
Total	8.8208e+13	2292	3.8485e+10	R-squared =	0.9974
				Adj R-squared =	0.9974
				Root MSE =	9989.7

y	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
x	-5.618982	24.07808	-0.23	0.815	-52.83611 41.59815
c.x#c.x	-30.70687	.9647104	-31.83	0.000	-32.59867 -28.81507
c.x#c.x#c.x	-15.02577	.0391987	-383.32	0.000	-15.10264 -14.9489
_cons	249.4734	309.3394	0.81	0.420	-357.1414 856.0882

. ovtest

```

Ramsey RESET test using powers of the fitted values of y
Ho: model has no omitted variables
      F(3, 2286) =      0.14
      Prob > F =      0.9390

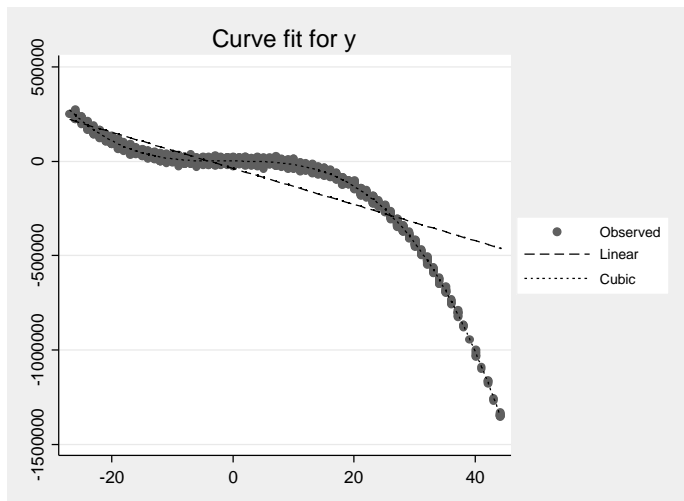
```

The original model had a pretty high value for R^2 . However, the scatterplot showed that the relationship between X and Y was not linear. Increases in X initially produced decreases in Y, then produced increases, then went back to producing decreases. This suggested a cubic model where X^2 and X^3 should be added. Once this was done the fit of the model was much better and the ovtest suggested that no other higher powers were needed. If she had stuck with the linear model, she would have alternated between underestimating and overestimating the expected value of Y given X:

```

. curvefit y x, f(1 5)
[Output deleted]

```



Appendix: Stata Code used in the exam

```

* Problem II
clear all
matrix input corr = (1,-.5,-.1,.6\-.5,1,.2,0\-.1,.2,1,0\-.6,0,0,1)
corr2data x1 x2 x3 x4, n(100) corr(corr)
* Confirm results
corr
reg x2 x1
reg x3 x2 x1
reg x4 x3 x2 x1
pathreg (x2 x1) (x3 x2 x1) (x4 x1 x2 x3)
sem (x2 <- x1) (x3 <- x2 x1) (x4 <- x1 x2 x3)
estat teffects
* Erroneous model
reg x4 x2

* Part III - Interaction effects
* Generate the variables by manipulating nhanes2f

```

```

* The manipulations produce the kind of relationships desired for the problem!
clear all
webuse nhanes2f, clear
keep health weight female
keep if !missing(health, weight, female)
set seed 123456
sample 2000, count
*gen older = female
replace weight = weight + (17 * female)
center weight, gen(offensive)
label variable offensive "How offensive is yikyak"
gen yikyak = (health-1) * 25 - .6*offensive - 13*female
label variable yikyak "How likely to continue using yikyak"
gen femXoffensive = female * offensive
* Do analyses
ttest yikyak, by(female)
nestreg: reg yikyak offensive female femXoffensive
ttest offensive, by(female)
* Additional analysis. This will plot the relationships
* and show differences in effects between females and males
quietly reg yikyak offensive i.female i.female#c.offensive
mcp offensive female, var1(20)

* Part IV-1: Piecewise regression
* Manipulate data. By construction, the effect of education is very different
* for lower grades than it is for higher.
clear all
use "http://www3.nd.edu/~rwilliam/statafiles/blwh.dta", clear
set seed 123456
replace educ = educ + rnormal()
gen inc = 2 * educ if educ <=12
replace inc = 8 * educ - 72 if educ > 12
replace inc = inc + rnormal(0, 6)
gen ses = educ * 3
gen liberalism = inc * 1.5
* Do analysis
reg liberalism ses
predict linear
label variable linear "linear"
scatter liberalism ses || line linear ses, scheme(sj) sort
mkspline seslow 36 seshi = ses
reg liberalism seslow seshi
predict spline
label variable spline "spline"
scatter liberalism ses || line spline ses, scheme(sj) sort

* IV-2 - Nonlinear relationships
*** Set up data
use "http://www.indiana.edu/~jlsoc/stata/spex_data/ordwarm2.dta", clear
corr2data e, sd(10000)
sum age
gen x = age - r(mean)
gen y = x - (30 * x^2) - (15* x^3) + e
*** Do analyses
reg y x
scatter y x, scheme(sj)
reg y x c.x#c.x c.x#c.x#c.x
ovtest
curvefit y x, f(1 5)

```